

A new parameter estimation method for the extended power Lindley distribution based on order statistics, with application

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Abstract

In this paper, we propose inference procedures for the estimation of parameters by using order statistics. First, we derive some new expressions for single and product moments of the order statistics from the extended power Lindley distribution. We then use these moments to obtain the best linear unbiased estimates (BLUEs) of the location and scale parameters based on Type-II right-censored samples. A real data set is analysed to illustrate the flexibility and importance of the model.

Key words: extended power Lindley distribution, order statistics, moments, best linear unbiased estimator.

1. Introduction

The Lindley distribution was proposed by Lindley (1958) in the context of fiducial and Bayesian statistics, which can be seen as a mixture of $\exp(\xi)$ and $\text{gamma}(2, \xi)$ distributions. Later, Ghitany et al. (2008) studied the statistical and mathematical properties of this distribution and also showed that this distribution performs better than the well-known exponential distribution in many ways. The Lindley distribution has only one scale parameter and is capable of modelling data with monotonic increasing failure rates and due to this, the Lindley distribution does not provide enough flexibility for analysing different types of lifetime data. To increase the flexibility for modeling purposes it will be useful to consider further alternatives of this distribution. Also, Bouchahed and Zeghdoudi (2018) and Zeghdoudi et al. (2018) generalized the Lindley distribution. Recently, the three parameter extended power Lindley distribution was proposed by Alkarni (2015) for the flexibility of purpose. The extended power Lindley (EPL) distribution is specified by the following probability density function (pdf):

$$f(x; \tau, \xi, \delta) = \frac{\tau \xi^2}{\xi + \delta} (1 + \delta x^\tau) x^{\tau-1} e^{-\xi x^\tau}, \quad x > 0; \quad \tau > 0, \xi > 0, \delta > 0 \quad (1)$$

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and the cumulative distribution function (cdf) is

$$F(x; \tau, \xi, \delta) = 1 - \left(1 + \frac{\delta \xi}{\xi + \delta} x^\tau\right) e^{-\xi x^\tau}, \quad x > 0; \quad \tau > 0, \xi > 0, \delta > 0. \quad (2)$$

For $\delta = 1$ and $\delta = 1$, $\tau = 1$, the EPL distribution reduces to the power Lindley (PL) and Lindley distributions respectively.

Let us consider the pdf of the scale-parameter EPL distribution as

$$f(x; \tau, \delta, \xi, \sigma) = \frac{\tau \xi^2}{\sigma(\xi + \delta)} \left[1 + \delta \left(\frac{x}{\sigma}\right)^\tau\right] \left(\frac{x}{\sigma}\right)^{\tau-1} e^{-\xi \left(\frac{x}{\sigma}\right)^\tau}, \quad x > 0; \quad \tau > 0, \delta > 0, \xi > 0, \sigma > 0, \quad (3)$$

and the pdf of the location-scale of the EPL distribution is

$$\begin{aligned} f(x; \tau, \delta, \xi, \sigma) &= \frac{\tau \xi^2}{\sigma(\xi + \delta)} \left[1 + \delta \left(\frac{x - \mu}{\sigma}\right)^\tau\right] \left(\frac{x - \mu}{\sigma}\right)^{\tau-1} \\ &\times e^{-\xi \left(\frac{x - \mu}{\sigma}\right)^\tau}, \quad x > \mu; \quad \tau > 0, \delta > 0, \xi > 0, \sigma > 0, \mu > 0, \quad (4) \end{aligned}$$

Order statistics play an important role in a wide range of theoretical and practical problems such as estimation of the problems, reliability analysis, quality control and strength of materials, for example, Balakrishnan and Cohan (1991), Sammel and Thomas (1997), Balakrishnan and Chandramouleeswaran, (1996), Sultan et al. (2000), Sultan and Balakrishnan (2000), Mahmoud et al. (2005), Jabeen et al. (2013), Sultan and AL-Thubyani (2016), Kumar et al. (2017), Kumar and Dey (2017a, 2017b), Ahsanullah and Alzaatreh (2018), Kumar and Goyal (2019a, 2019b), Kumar et al. (2020a, 2020b) and the references cited therein.

The rest of the paper is organized as follows. In Section 2, we introduce some lemmas on the EPL distribution and we discuss some expressions for single and double moments of order statistics. We use these moments to obtain the BLUEs for μ and σ in Section 3. In Section 4, we performed a numerical study using R software and we see that these expressions provide precise numerical evaluations. A real data application is presented in Section 5. Finally, we conclude the paper in Section 6.

2. Moments of Order Statistics

In this section, we provide and prove some lemmas and some new expressions for single and product moments of order statistics for the given a random sample X_1, X_2, \dots, X_n from the EPL distribution. Let $X_{(1)} \leq X_{(2)} \leq \dots \leq X_{(n)}$ denote the order statistics corresponding to X_1, X_2, \dots, X_n . The pdf of the r th order statistic is given by Arnold et al. (2003) and David and Nagaraja (2003) as

$$f_{X_{(r:n)}}(x) = C_{r:n} F^{r-1}(x) [1 - F(x)]^{n-r} f(x) \quad (5)$$

for $x > 0$. The joint pdf of the r th and s th order statistics is given by Arnold et al. (2003) and David and Nagaraja (2003) as

$$f_{X_{(r:n)}, X_{(s:n)}}(x, y) = C_{r,s:n} F^{r-1}(x) [F(y) - F(x)]^{s-1-r} [1 - F(x)]^{n-s} f(x)f(y) \quad (6)$$

for $0 < x < y$, where $C_{r:n} = \frac{n!}{(r-1)!(n-r)!}$ and $C_{r,s:n} = \frac{n!}{(r-1)!(s-r-1)!(n-s)!}$.

2.1. Some Lemmas

Here, we provide and prove two lemmas.

Lemma 1 Let $f(x)$ and $F(x)$ be given by (1) and (2), respectively. For $a > 0$, $b > 0$ and $p > 0$, let

$$K(a, b, p) = \int_0^\infty x^p F^a(x) [1 - F(x)]^b f(x) dx.$$

Then,

$$K(a, b, p) = \frac{\xi^2}{(\xi + \delta)^{b+1}} \sum_{i_1=0}^a \sum_{i_2=0}^{i_1+b} \sum_{i_3=0}^{i_2+1} (-1)^{i_1} \frac{\xi^{i_2} \delta^{b+i_1-i_2+i_3}}{(\xi + \delta)^{i_1}} \binom{a}{i_1} \binom{i_1+b}{i_2} \binom{i_2+1}{i_3} \times \frac{\Gamma\left(\frac{p+\tau(i_3+1)}{\tau}\right)}{[\xi(i_1+b+1)]^{\frac{p+\tau(i_3+1)}{\tau}}}.$$

Proof. By using binomial expansion in (5) and then from (1) and (2), we get

$$\begin{aligned} K(a, b, p) &= \sum_{i_1=0}^a (-1)^{i_1} \binom{a}{i_1} \int_0^\infty x^p [1 - F(x)]^{i_1+b} f(x) dx. \\ &= \frac{\tau \xi^2}{(\xi + \delta)} \sum_{i_1=0}^a (-1)^{i_1} \binom{a}{i_1} \int_0^\infty x^{p+\tau-1} (1 + \delta x^\tau)^{e^{-\xi(i_1+b+1)x^\tau}} \\ &\times \left[\left(1 + \frac{\delta \xi}{\xi + \delta} x^\tau \right) \right]^{i_1+b} dx \\ &= \frac{\tau \xi^2}{(\xi + \delta)^{b+1}} \sum_{i_1=0}^a \sum_{i_2=0}^{i_1+b} \sum_{i_3=0}^{i_2+1} (-1)^{i_1} \frac{\xi^{i_2} \delta^{b+i_1-i_2+i_3}}{(\xi + \delta)^{i_1}} \binom{a}{i_1} \binom{i_1+b}{i_2} \binom{i_2+1}{i_3} \\ &\times \int_0^\infty x^{p+\tau(i_3+1)-1} e^{-\xi(i_1+b+1)x^\tau} dx \\ &= \frac{\xi^2}{(\xi + \delta)^{b+1}} \sum_{i_1=0}^a \sum_{i_2=0}^{i_1+b} \sum_{i_3=0}^{i_2+1} (-1)^{i_1} \frac{\xi^{i_2} \delta^{b+i_1-i_2+i_3}}{(\xi + \delta)^{i_1}} \binom{a}{i_1} \binom{i_1+b}{i_2} \binom{i_2+1}{i_3} \\ &\times \frac{1}{[\xi(i_1+b+1)]^{\frac{p+\tau(i_3+1)}{\tau}}} \int_0^\infty z^{\frac{p+\tau i_3}{\tau}} e^{-z} dz, \end{aligned}$$

where $z = \xi(i_1 + b + 1)x^\tau$.

Lemma 2 Let $f(x)$ and $F(x)$ be given by (1) and (2), respectively. For $a > 0, b > 0, c, p > 0$ and $q > 0$, let

$$L(a, b, c, p, q) = \int_0^\infty \int_x^\infty x^p y^q [F(x)]^a [F(y) - F(x)]^b [1 - F(y)]^c f(x) f(y) dy dx.$$

Then,

$$\begin{aligned} L(a, b, c, p, q) &= \frac{\xi^4}{(\xi + \delta)^{b+c+2}} \sum_{i_1=0}^a \sum_{i_2=0}^b \sum_{i_3=0}^{i_1+b-i_2} \sum_{i_4=0}^{i_2+c} \sum_{i_5=0}^{i_3+1} \sum_{i_6=0}^{i_4+1} \sum_{i_7=0}^{\frac{q+\tau i_6}{\tau}} (-1)^{i_1+i_2} \binom{a}{i_1} \binom{b}{i_2} \\ &\times \binom{i_1+b-i_2}{i_3} \binom{i_2+c}{i_4} \binom{i_3+1}{i_5} \binom{i_4+1}{i_6} \frac{\xi^{i_3+i_4} \delta^{i_1+b-i_3-i_4+i_5+i_6}}{(\xi + \delta)^{i_1}} \\ &\times \frac{\left(\frac{q+\tau i_6}{\tau}\right)!}{[\xi(i_2+c+1)]^{\frac{q+\tau(i_6+1)}{\tau}}} \frac{[\xi(i_2+c+1)]^{i_7}}{i_7!} \frac{\Gamma\left(\frac{p+\tau(i_5+i_7+1)}{\tau}\right)}{[\xi(i_1+b+c+2)]^{\frac{p+\tau(i_5+i_7+1)}{\tau}}}. \end{aligned}$$

Proof. By using binomial expansion in (6) and then from (1) and (2), we get

$$\begin{aligned} L(a, b, c, p, q) &= \sum_{i_1=0}^a \sum_{i_2=0}^b (-1)^{i_1+i_2} \binom{a}{i_1} \binom{b}{i_2} \\ &\times \int_0^\infty \int_x^\infty x^p y^q [1 - F(x)]^{i_1+b-i_2} [1 - F(y)]^{i_2+c} f(x) f(y) dy dx \\ &= \frac{\tau^2 \xi^4}{(\xi + \delta)^2} \sum_{i_1=0}^a \sum_{i_2=0}^b (-1)^{i_1+i_2} \binom{a}{i_1} \binom{b}{i_2} \int_0^\infty \int_x^\infty x^{p+\tau-1} y^{q+\tau-1} \\ &\times (1 + \delta x^\tau)(1 + \delta y^\tau) e^{-\xi(i_1+b-i_2+1)x^\tau} e^{-\xi(i_2+c+1)y^\tau} \\ &\times \left[\left(1 + \frac{\xi \delta}{\xi + \delta} x^\tau\right) \right]^{i_1+b-i_2} \left[\left(1 + \frac{\xi \delta}{\xi + \delta} y^\tau\right) \right]^{i_2+c} dx dy \\ &= \frac{\tau^2 \xi^4}{(\xi + \delta)^{b+c+2}} \sum_{i_1=0}^a \sum_{i_2=0}^b \sum_{i_3=0}^{i_1+b-i_2} \sum_{i_4=0}^{i_2+c} \sum_{i_5=0}^{i_3+1} \sum_{i_6=0}^{i_4+1} (-1)^{i_1+i_2} \\ &\times \frac{\xi^{i_3+i_4} \delta^{i_1+b-i_3-i_4+i_5+i_6}}{(\xi + \delta)^{i_1}} \binom{a}{i_1} \binom{b}{i_2} \binom{i_1+b-i_2}{i_3} \binom{i_2+c}{i_4} \binom{i_3+1}{i_5} \\ &\times \binom{i_4+1}{i_6} \frac{1}{[\xi(i_2+c+1)]^{\frac{q+\tau(i_6+1)}{\tau}}} \int_0^\infty x^{p+\tau(i_5+1)-1} e^{-\xi(i_1+b-i_2+1)x^\tau} \\ &\times \Gamma\left(\frac{q+\tau(i_6+1)}{\tau}, \xi(i_2+c+1)x^\tau\right) dx, \end{aligned}$$

By using the relation

$$\Gamma(p, y) = (p - 1)! e^{-y} \sum_{l=0}^{p-1} \frac{y^l}{l!}.$$

$$\begin{aligned}
 L(a, b, c, p, q) &= \frac{\tau^2 \xi^4}{(\xi + \delta)^{b+c+2}} \sum_{i_1=0}^a \sum_{i_2=0}^b \sum_{i_3=0}^{i_1+b-i_2} \sum_{i_4=0}^{i_2+c} \sum_{i_5=0}^{i_3+1} \sum_{i_6=0}^{i_4+1} \sum_{i_7=0}^{\frac{q+\tau i_6}{\tau}} (-1)^{i_1+i_2} \binom{a}{i_1} \binom{b}{i_2} \\
 &\times \frac{\xi^{i_3+i_4} \delta^{i_1+b-i_3-i_4+i_5+6}}{(\xi + \delta)^{i_1}} \binom{i_1+b-i_2}{i_3} \binom{i_2+c}{i_4} \binom{i_3+1}{i_5} \binom{i_4+1}{i_6} \\
 &\times \frac{\left(\frac{q+\tau i_6}{\tau}\right)!}{[\xi(i_2+c+1)]^{\frac{q+\tau(i_6+1)}{\tau}}} \frac{[\xi(i_2+c+1)]^{i_7}}{i_7!} \\
 &\times \int_0^\infty x^{p+\tau(i_5+i_7+1)-1} e^{-\xi(i_1+b+c+2)x^\tau} dx \\
 &= \frac{\xi^4}{(\xi + \delta)^{b+c+2}} \sum_{i_1=0}^a \sum_{i_2=0}^b \sum_{i_3=0}^{i_1+b-i_2} \sum_{i_4=0}^{i_2+c} \sum_{i_5=0}^{i_3+1} \sum_{i_6=0}^{i_4+1} \sum_{i_7=0}^{\frac{q+\tau i_6}{\tau}} (-1)^{i_1+i_2} \binom{a}{i_1} \binom{b}{i_2} \\
 &\times \frac{\xi^{i_3+i_4} \delta^{i_1+b-i_3-i_4+i_5+i_6}}{(\xi + \delta)^{i_1}} \binom{i_1+b-i_2}{i_3} \binom{i_2+c}{i_4} \binom{i_3+1}{i_5} \binom{i_4+1}{i_6} \\
 &\times \frac{\left(\frac{q+\tau i_6}{\tau}\right)!}{[\xi(i_2+c+1)]^{\frac{q+\tau(i_6+1)}{\tau}}} \frac{[\xi(i_2+c+1)]^{i_7}}{i_7!} \\
 &\times \frac{1}{[\xi(i_1+b+c+2)]^{\frac{p+\tau(i_5+i_7+1)}{\tau}}} \int_0^\infty z^{\frac{p+\tau(i_5+i_7)}{\tau}} e^{-z} dz,
 \end{aligned}$$

where $z = \xi(i_1 + b + c + 2)x^\tau$.

2.2. Single moments

Here, we present expressions for the single moments of r th order statistics, $E(X_{r:n}^{(p)}) = \mu_{r:n}^{(p)}$ from the EPL distribution because these are very important to calculate the variance and draw the inferential techniques.

Theorem 1. For the EPL distribution given in (1) and for, $1 \leq r \leq n$

$$\begin{aligned}
 \mu_{r:n}^{(p)} &= \frac{\xi^2 C_{r:n}}{(\xi + \delta)^{n-r+1}} \sum_{i_1=0}^{r-1} \sum_{i_2=0}^{i_1+n-r-i_2+1} \sum_{i_3=0}^{i_2} (-1)^{i_1} \binom{r-1}{i_1} \binom{i_1+n-r}{i_2} \binom{i_2+1}{i_3} \\
 &\times \frac{\xi^{i_2} \delta^{i_1+n-r-i_2+i_3}}{(\xi + \delta)^{i_1}} \frac{\Gamma\left(\frac{p+\tau(i_3+1)}{\tau}\right)}{[\xi(i_1+n-r+1)]^{\frac{p+\tau(i_3+1)}{\tau}}}. \tag{7}
 \end{aligned}$$

Proof. By using (5), we get

$$\mu_{r:n}^{(p)} = C_{r:n} \int_0^\infty x^p F^{r-1}(x) [1 - F(x)]^{n-r} f(x) dx. \tag{8}$$

The result follows from applying lemma 1.

Special Cases

i For $\tau = 1$ and $\delta = 1$ in (7), we obtain relation for order statistic of the Lindley distribution

$$\begin{aligned} \mu_{r:n}^{(p)} &= \frac{\xi^2 C_{r:n}}{(\xi + 1)^{n-r+1}} \sum_{i_1=0}^{r-1} \sum_{i_2=0}^{i_1+n-r-i_2+1} \sum_{i_3=0}^{i_2+1} (-1)^{i_1} \binom{r-1}{i_1} \binom{i_1+n-r}{i_2} \binom{i_2+1}{i_3} \\ &\times \frac{\xi^{i_2}}{(\xi + 1)^{i_1}} \frac{\Gamma(p + i_3 + 1)}{[\xi(i_1 + n - r + 1)]^{p+i_3+1}}, \end{aligned}$$

as obtained by Sultan and AL-Thubyani (2016).

ii For $\delta = 1$ in (7), we obtain relation for order statistic of the power Lindley distribution

$$\begin{aligned} \mu_{r:n}^{(p)} &= \frac{\xi^2 C_{r:n}}{(\xi + 1)^{n-r+1}} \sum_{i_1=0}^{r-1} \sum_{i_2=0}^{i_1+n-r-i_2+1} \sum_{i_3=0}^{i_2+1} (-1)^{i_1} \binom{r-1}{i_1} \binom{i_1+n-r}{i_2} \binom{i_2+1}{i_3} \\ &\times \frac{\xi^{i_2}}{(\xi + 1)^{i_1}} \frac{\Gamma(\frac{p+\tau(i_3+1)}{\tau})}{[\xi(i_1 + n - r + 1)]^{\frac{p+\tau(i_3+1)}{\tau}}}, \end{aligned}$$

as obtained by Kumar and Goyal (2019a).

iii If $p = n = r = 1$ in (7), we unordered mean of the extended power Lindley distribution

$$\mu_{1:1}^{(1)} = \frac{\xi^2}{(\xi + \delta)} \sum_{i_3=0}^1 \frac{\delta^{i_3} \Gamma(\frac{\tau(i_3+1)+1}{\tau})}{\xi^{\frac{\tau(i_3+1)+1}{\tau}}} = E(X),$$

as obtained by Alkarni (2015).

In particular, the first moment (mean) of the r -th order statistic is

$$\mu_{r:n} = \mu_{r:n}^{(1)} = \vartheta(\tau, \xi, \delta, r, n, 1),$$

where

$$\begin{aligned} \vartheta(\tau, \xi, \delta, r, n, k) &= \frac{\xi^2 C_{r:n}}{(\xi + \delta)^{n-r+1}} \sum_{i_1=0}^{r-1} \sum_{i_2=0}^{i_1+n-r-i_2+1} \sum_{i_3=0}^{i_2+1} (-1)^{i_1} \binom{r-1}{i_1} \binom{i_1+n-r}{i_2} \binom{i_2+1}{i_3} \\ &\times \frac{\xi^{i_2} \delta^{i_1+n-r-i_2+i_3}}{(\xi + \delta)^{i_1}} \frac{\Gamma(\frac{k+\tau(i_3+1)}{\tau})}{[\xi(i_1 + n - r + 1)]^{\frac{k+\tau(i_3+1)}{\tau}}}. \end{aligned}$$

In addition, the variance of $X_{r:n}$ is found to be

$$\sigma_{r:n}^2 = \mu_{r:n}^{(2)} - [\mu_{r:n}^{(1)}]^2 = \vartheta(\tau, \xi, \delta, r, n, 2) - [\vartheta(\tau, \xi, \delta, r, n, 1)]^2.$$

2.3. Double moments

Here, we provide relation for the double moment of r th and s th order statistics, $E\left(X_{r,s:n}^{(p,q)}\right) = \mu_{r,s:n}^{(p,q)}$ from the EPL distribution, which are useful to calculate the variances, covariances, BLUEs and other inferential techniques.

Theorem 2. For the EPL distribution given in (1) and for $1 \leq r < s \leq n$

$$\begin{aligned} \mu_{r,s:n}^{(p,q)} &= C_{r,s:n} \frac{\xi^4}{(\xi + \delta)^{n-r+1}} \sum_{i_1=0}^{r-1} \sum_{i_2=0}^{s-r-1} \sum_{i_3=0}^{i_1+s-r-i_2-1} \sum_{i_4=0}^{i_2+n-s-i_3+1} \sum_{i_5=0}^{i_4+1} \sum_{i_6=0}^{q+i_6} \sum_{i_7=0}^{\frac{q+i_6}{\tau}} (-1)^{i_1+i_2} \\ &\times \frac{\xi^{i_3+i_4} \delta^{i_1+s-r-1-i_3-i_4+i_5+i_6}}{(\xi + \delta)^{i_1}} \binom{r-1}{i_1} \binom{s-r-1}{i_2} \binom{i_1+s-r-i_2-1}{i_3} \\ &\times \binom{i_2+n-s}{i_4} \binom{i_3+1}{i_5} \binom{i_4+1}{i_6} \frac{\left(\frac{q+i_6}{\tau}\right)!}{[\xi(i_2+n-s+1)]^{\frac{q+i_6}{\tau}}} \\ &\times \frac{[\xi(i_2+n-s+1)]^{i_7}}{i_7!} \frac{\Gamma\left(\frac{p+\tau(i_5+i_7+1)}{\tau}\right)}{[\xi(i_1+n-r+1)]^{\frac{p+\tau(i_5+i_7+1)}{\tau}}}. \end{aligned}$$

Proof. By using (6), we obtain

$$\mu_{r,s:n}^{(p,q)} = C_{r,s:n} \int_0^\infty \int_x^\infty x^p y^q F^{r-1}(x) [F(y) - F(x)]^{s-1-r} [1 - F(y)]^{n-s} f(x)f(y) dx dy \quad (9)$$

The result follows from applying lemma 2.

Special Cases

i For $\tau = 1$ and $\delta = 1$ in Theorem 2, we obtain the relation for order statistic of the Lindley distribution

$$\begin{aligned} \mu_{r,s:n}^{(p,q)} &= \frac{\xi^4 C_{r,s:n}}{(\xi + 1)^{n-r+1}} \sum_{i_1=0}^{r-1} \sum_{i_2=0}^{s-r-1} \sum_{i_3=0}^{i_1+s-r-i_2-1} \sum_{i_4=0}^{i_2+n-s-i_3+1} \sum_{i_5=0}^{i_4+1} \sum_{i_6=0}^{q+i_6} \sum_{i_7=0}^{\frac{q+i_6}{\tau}} (-1)^{i_1+i_2} \frac{\xi^{i_3+i_4}}{(\xi + 1)^{i_1}} \\ &\times \binom{r-1}{i_1} \binom{s-r-1}{i_2} \binom{i_1+s-r-i_2-1}{i_3} \binom{i_2+n-s}{i_4} \binom{i_3+1}{i_5} \binom{i_4+1}{i_6} \\ &\times \frac{(q+i_6)!}{[\xi(i_2+n-s+1)]^{q+i_6+1}} \frac{[\xi(i_2+n-s+1)]^{i_7}}{i_7!} \frac{\Gamma(p+i_5+i_7+1)}{[\xi(i_1+n-r+1)]^{p+i_5+i_7+1}}. \end{aligned}$$

as obtained by Sultan and AL-Thubyani (2016).

ii For $\delta = 1$ in Theorem 2, we get the explicit expression for order statistic of the power Lindley distribution

$$\begin{aligned} \mu_{r,s;n}^{(p,q)} &= \frac{\xi^4 C_{r,s;n}}{(\xi + 1)^{n-r+1}} \sum_{i_1=0}^{r-1} \sum_{i_2=0}^{s-r-1} \sum_{i_3=0}^{i_1+s-r-i_2-1} \sum_{i_4=0}^{i_2+n-s} \sum_{i_5=0}^{i_3+1} \sum_{i_6=0}^{i_4+1} \sum_{i_7=0}^{q+\tau i_6} (-1)^{i_1+i_2} \frac{\xi^{i_3+i_4}}{(\xi + 1)^{i_1}} \\ &\times \binom{r-1}{i_1} \binom{s-r-1}{i_2} \binom{i_1+s-r-i_2-1}{i_3} \binom{i_2+n-s}{i_4} \binom{i_3+1}{i_5} \binom{i_4+1}{i_6} \\ &\times \frac{\left(\frac{q+\tau i_6}{\tau}\right)!}{[\xi(i_2+n-s+1)]^{\frac{q+\tau(i_6+1)}{\tau}}} \frac{[\xi(i_2+n-s+1)]^{i_7}}{i_7!} \frac{\Gamma\left(\frac{p+\tau(i_5+i_7+1)}{\tau}\right)}{[\xi(i_1+n-r+1)]^{\frac{p+\tau(i_5+i_7+1)}{\tau}}}. \end{aligned}$$

as obtained by Kumar and Goyal (2019a).

For simplicity, we denote the (1, 1)-th moment of $X_{r;n}$ and $X_{s;n}$, which are also called the simple product moment of these order statistics, by $\mu_{r,s;n}$. The simple product moments are used for evaluating the covariances, in other words

$$\sigma_{r,s;n} = Cov(X_{r;n}, X_{s;n}) = \mu_{r,s;n} - \mu_{r;n}\mu_{s;n}.$$

3. Location and scale parameter estimation based on Type-II right-censored sample

In this section we have obtained the BLUEs of the location and scale parameters to approximate population parameters. Let $Y_{1:n} \leq Y_{2:n} \leq \dots \leq Y_{(n-c):n}$, $c = 0, 1, \dots, n-1$ denote Type-II right-censored sample from the location-scale parameter extended power Lindley distribution in Equation (1). Let us denote $Z_{r;n} = (Y_{r;n} - \mu)/\sigma$, $E(Z_{r;n}) = \mu_{r;n}^{(1)}$, $1 \leq r \leq (n-c)$, and $Cov(Z_{r;n}, Z_{s;n}) = \sigma_{r,s;n} = \mu_{r,s;n}^{(1,1)} - \mu_{r;n}^{(1)}\mu_{s;n}^{(1)}$, $1 \leq r < s \leq (n-c)$. We shall use the following notations:

$$\mathbf{Y} = (Y_{1:n}, Y_{2:n}, \dots, Y_{(n-c):n})^T,$$

$$\boldsymbol{\mu} = (\mu_{1:n}, \mu_{2:n}, \dots, \mu_{(n-c):n})^T,$$

$$\mathbf{1} = \underbrace{(1, 1, \dots, 1)}_{n-c}^T,$$

and

$$\boldsymbol{\Sigma} = ((\sigma_{r,s})), \quad 1 \leq r, s \leq n-c.$$

The BLUEs of μ and σ are given by Arnold *et al.* (2003)

$$\mu^* = \left\{ \frac{\mu^T \Sigma^{-1} \mu \mathbf{1}^T \Sigma^{-1} - \mu^T \Sigma^{-1} \mathbf{1} \mu^T \Sigma^{-1}}{(\mu^T \Sigma^{-1} \mu)(\mathbf{1}^T \Sigma^{-1} \mathbf{1}) - (\mu^T \Sigma^{-1} \mathbf{1})^2} \right\} Y_{r:n} = \sum_{r=1}^{n-c} a_r Y_{r:n}, \tag{10}$$

$$\sigma^* = \left\{ \frac{\mathbf{1}^T \Sigma^{-1} \mathbf{1} \mu^T \Sigma^{-1} - \mathbf{1}^T \Sigma^{-1} \mu \mathbf{1}^T \Sigma^{-1}}{(\mu^T \Sigma^{-1} \mu)(\mathbf{1}^T \Sigma^{-1} \mathbf{1}) - (\mu^T \Sigma^{-1} \mathbf{1})^2} \right\} Y_{r:n} = \sum_{r=1}^{n-c} b_r Y_{r:n}. \tag{11}$$

Furthermore, the variances and covariance of these BLUEs are given by Arnold *et al.* (2003)

$$Var(\mu^*) = \sigma^2 \left\{ \frac{\mu^T \Sigma^{-1} \mu}{(\mu^T \Sigma^{-1} \mu)(\mathbf{1}^T \Sigma^{-1} \mathbf{1}) - (\mu^T \Sigma^{-1} \mathbf{1})^2} \right\} \sigma^2 V_1, \tag{12}$$

$$Var(\sigma^*) = \sigma^2 \left\{ \frac{\mathbf{1}^T \Sigma^{-1} \mathbf{1}}{(\mu^T \Sigma^{-1} \mu)(\mathbf{1}^T \Sigma^{-1} \mathbf{1}) - (\mu^T \Sigma^{-1} \mathbf{1})^2} \right\} \sigma^2 V_2, \tag{13}$$

and

$$Cov(\mu^*, \sigma^*) = \sigma^2 \left\{ \frac{-\mu^T \Sigma^{-1} \mathbf{1}}{(\mu^T \Sigma^{-1} \mu)(\mathbf{1}^T \Sigma^{-1} \mathbf{1}) - (\mu^T \Sigma^{-1} \mathbf{1})^2} \right\} \sigma^2 V_3. \tag{14}$$

4. Numerical Results

The relations obtained in the preceding sections allow us to evaluate the expected values, second moments, variances, product moment and covariances of order statistics from samples of sizes up to 10 for various values of the parameters. The relation in (7) can be used to compute the expected values, second order moments and variances of all order statistics for sample sizes $n = 1, 2, 3, \dots, 10$. In Tables 1, we have presented expected values, second order moments, variances, skewness and kurtosis of the r th order statistics from the EPL distribution for $n = 1, 2, 3, \dots, 10$ and $\tau = 2, \delta = 0.5$ and $\xi = 5$. One can see that the means and variances are decreasing both with respect to n and ξ . In Table 2 we have reported the double moments and covariances of the r th and s th order statistics from the EPL distribution for $n = 1, 2, 3, \dots, 10$ and $\tau = 2, \delta = 0.5$ and $\xi = 5, 10$. From Table 3, one can observe that product moments are decreasing with respect to n .

The coefficients of the BLUEs for Type-II right-censored samples of sample sizes $n = 7, 10, \tau = 2, \delta = 0.5$ and $\xi = 5$ and different censoring cases $c = 0(1)([n/2] - 1)$ are displayed in Tables 3 and 4. Also, the variances and covariances of the BLUEs are provided in Table 5.

Table 1: Moments, variances skewness and kurtosis of order statistic from EPL distribution $\tau = 2, \delta = 0.5$ and $\xi = 5$

r	n	$E(X)$	$E(X^2)$	$V(X)$	δ_1	δ_2	γ_1	γ_2		
1	1	0.414348	0.218182	0.046498	0.379571	3.207079	0.616093	0.207079		
	2	0.293422	0.109504	0.023408	0.387238	3.219388	0.622284	0.219388		
	3	0.239709	0.073106	0.015646	0.389946	3.229529	0.624457	0.229529		
	4	0.207652	0.054870	0.011751	0.391543	3.234472	0.625734	0.234472		
	5	0.185762	0.043916	0.009408	0.392307	3.236857	0.626344	0.236857		
	6	0.169597	0.036608	0.007845	0.394587	3.229235	0.628161	0.229235		
	7	0.157029	0.031385	0.006727	0.395586	3.221588	0.628956	0.221588		
	8	0.146897	0.027467	0.005888	0.395741	3.229525	0.629079	0.229525		
	9	0.138503	0.024418	0.005235	0.394610	3.235330	0.628180	0.235330		
	10	0.131400	0.021978	0.004712	0.396862	3.226521	0.629970	0.226521		
2	2	0.535273	0.326860	0.040343	0.241593	3.215795	0.491521	0.215795		
	3	0.400850	0.182301	0.021620	0.189444	3.115636	0.435251	0.115636		
	4	0.335877	0.127814	0.015001	0.174957	3.088871	0.418279	0.088871		
	5	0.295213	0.098686	0.011535	0.170039	3.070969	0.412358	0.070969		
	6	0.266589	0.080456	0.009386	0.167111	3.066828	0.408792	0.066828		
	7	0.245001	0.067944	0.007919	0.165394	3.065728	0.406687	0.065728		
	8	0.227956	0.058815	0.006851	0.164916	3.061962	0.406099	0.061962		
	9	0.214051	0.051856	0.006038	0.164358	3.069852	0.405411	0.069852		
	10	0.202423	0.046373	0.005398	0.164208	3.048848	0.405226	0.048848		
	3	3	0.602485	0.399139	0.036151	0.225634	3.273408	0.475009	0.273408	
4		0.465823	0.236788	0.019797	0.145361	3.125628	0.381262	0.125628		
5		0.396873	0.171505	0.013997	0.123096	3.072745	0.350850	0.072745		
6		0.352462	0.135146	0.010917	0.112742	3.054180	0.355770	0.054180		
7		0.320559	0.111736	0.008978	0.108337	3.051939	0.329145	0.051939		
8		0.296135	0.095332	0.007636	0.105760	3.035578	0.325207	0.035578		
9		0.276626	0.083171	0.006649	0.105379	3.008423	0.324621	0.008423		
10		0.260564	0.073785	0.005891	0.104265	3.025598	0.322900	0.025598		
4		4	0.648039	0.453256	0.033301	0.229320	3.321976	0.478874	0.321976	
		5	0.511790	0.280311	0.018382	0.130464	3.147984	0.361198	0.147984	
	6	0.441284	0.207865	0.013133	0.102281	3.090069	0.319815	0.090069		
	7	0.394999	0.166359	0.010335	0.090404	3.056551	0.300673	0.056551		
	8	0.361267	0.139076	0.008562	0.084264	3.040724	0.290282	0.040724		
	9	0.335151	0.119652	0.007326	0.081505	3.010495	0.285491	0.010495		
	10	0.314105	0.105073	0.006411	0.078213	3.023921	0.279667	0.023921		
	5	5	0.682101	0.496492	0.031230	0.237912	3.362265	0.487762	0.362265	
		6	0.547043	0.316534	0.017278	0.124926	3.165975	0.353448	0.165975	
		7	0.475998	0.238994	0.012420	0.092794	3.093716	0.304621	0.093716	
8		0.428730	0.193642	0.009833	0.078295	3.075215	0.279812	0.075215		
9		0.393912	0.163355	0.008188	0.071074	3.053238	0.266597	0.053238		
10		0.366721	0.141522	0.007038	0.067297	3.019226	0.259417	0.019226		
6		6	0.709112	0.532483	0.029643	0.247927	3.393989	0.497922	0.393989	
		7	0.575461	0.347549	0.016394	0.123722	3.179939	0.351742	0.179939	
		8	0.504359	0.266206	0.011828	0.087950	3.098604	0.296564	0.098604	
		9	0.456584	0.217871	0.009402	0.072183	3.067347	0.268669	0.067347	
	10	0.421103	0.185187	0.007859	0.063867	3.048346	0.252719	0.048346		
	7	7	0.731387	0.563306	0.028379	0.257507	3.423281	0.507451	0.423281	
		8	0.599162	0.374664	0.015668	0.123781	3.197735	0.351825	0.197735	
		9	0.528246	0.290374	0.011330	0.081133	3.116019	0.284838	0.116019	
		10	0.480238	0.239660	0.009031	0.067515	3.077528	0.259836	0.077528	
		8	8	0.750277	0.590255	0.027339	0.267491	3.444852	0.517195	0.444852
9			0.619424	0.398746	0.015060	0.125692	3.209343	0.354531	0.209343	
10			0.548821	0.312108	0.010904	0.083717	3.116898	0.289339	0.116898	
9			9	0.766633	0.614193	0.026467	0.276388	3.468726	0.525726	0.468726
			10	0.637074	0.420406	0.014543	0.127287	3.217234	0.356773	0.217234
			10	0.781029	0.635725	0.025719	0.285348	3.482843	0.534180	0.482843

Table 2: Covariances of order statistics

		$\tau = 2, \xi = 3, \delta = 0.5$					$\tau = 2, \xi = 10, \delta = 0.5$					
s	r	n	$\mu_{r,s:n}$	$E(X_{r:n})$	$E(X_{s:n})$	$\sigma_{r,s:n}$	$\mu_{r,s:n}$	$E(X_{r:n})$	$\mu_{r,s:n}$	$\sigma_{r,s:n}$		
2	1	2	0.125100	0.293422	0.535273	-0.03196	0.059192	0.202969	0.370875	-0.01608		
		3	0.073220	0.239709	0.400850	-0.02287	0.034402	0.165748	0.277412	-0.01158		
		4	0.052133	0.207652	0.335877	-0.01761	0.024407	0.143552	0.232333	-0.00895		
		5	0.040552	0.185762	0.295213	-0.01429	0.018943	0.128403	0.204150	-0.00727		
		6	0.033203	0.169597	0.266589	-0.01201	0.015488	0.117219	0.184323	-0.00612		
		7	0.028119	0.157029	0.245001	-0.01035	0.013102	0.108526	0.169376	-0.00528		
		8	0.024389	0.146897	0.227956	-0.00910	0.011355	0.101518	0.157579	-0.00464		
		9	0.021534	0.138503	0.214051	-0.00811	0.010020	0.095714	0.147956	-0.00414		
		10	0.019279	0.131400	0.202423	-0.00732	0.008966	0.090803	0.139911	-0.00374		
		3	1	3	0.146966	0.239709	0.602485	0.002545	0.066242	0.165748	0.417607	-0.00298
4	0.100791			0.207652	0.465823	0.004062	0.044251	0.143552	0.322490	-0.00204		
5	0.079630			0.185762	0.396873	0.005906	0.034146	0.128403	0.274609	-0.00112		
6	0.067164			0.169597	0.352462	0.007387	0.028186	0.117219	0.243803	-0.00039		
7	0.058882			0.157029	0.320559	0.008545	0.024224	0.108526	0.221690	0.000165		
8	0.052962			0.146897	0.296135	0.009461	0.021392	0.101518	0.204769	0.000604		
9	0.048512			0.138503	0.276626	0.010199	0.019262	0.095714	0.191258	0.000956		
10	0.045043			0.131400	0.260564	0.010805	0.017601	0.090803	0.180137	0.001244		
3	2			3	0.155112	0.400850	0.602485	-0.08639	0.076932	0.277412	0.417607	-0.03892
4	0.087823			0.335877	0.465823	-0.06864	0.044543	0.232333	0.322490	-0.03038		
5	0.059381	0.295213	0.396873	-0.05778	0.031058	0.204150	0.274609	-0.02500				
6	0.043331	0.266589	0.352462	-0.05063	0.023522	0.184323	0.243803	-0.02142				
7	0.032948	0.245001	0.320559	-0.04559	0.018679	0.169376	0.221690	-0.01887				
8	0.025656	0.227956	0.296135	-0.04185	0.015295	0.157579	0.204769	-0.01697				
9	0.020245	0.214051	0.276626	-0.03897	0.012794	0.147956	0.191258	-0.01550				
10	0.016067	0.202423	0.260564	-0.03668	0.010868	0.139911	0.180137	-0.01434				
4	1	4	0.157123	0.207652	0.648039	0.022556	0.067875	0.143552	0.449313	0.003375		
		5	0.116281	0.185762	0.511790	0.021210	0.048633	0.128403	0.354410	0.003126		
		6	0.096719	0.169597	0.441284	0.021879	0.039214	0.117219	0.305414	0.003413		
		7	0.084955	0.157029	0.394999	0.022929	0.033425	0.108526	0.273287	0.003766		
		8	0.077118	0.146897	0.361267	0.024049	0.029472	0.101518	0.249893	0.004104		
		9	0.071578	0.138503	0.335151	0.025159	0.026598	0.095714	0.231790	0.004413		
		10	0.067511	0.131400	0.314105	0.026238	0.024416	0.090803	0.217207	0.004693		
		2	4	0.172829	0.335877	0.648039	-0.04483	0.084966	0.232333	0.449313	-0.01942	
		5	0.112133	0.295213	0.511790	-0.03895	0.055696	0.204150	0.354410	-0.01666		
		6	0.082851	0.266589	0.441284	-0.03479	0.041888	0.184323	0.305414	-0.01441		
7	0.064709	0.245001	0.394999	-0.03207	0.033554	0.169376	0.273287	-0.01273				
8	0.052012	0.227956	0.361267	-0.03034	0.027892	0.157579	0.249893	-0.01149				
9	0.042428	0.214051	0.335151	-0.02931	0.023756	0.147956	0.231790	-0.01054				
10	0.034802	0.202423	0.314105	-0.02878	0.020579	0.139911	0.217207	-0.00981				
3	2	4	0.179898	0.465823	0.648039	-0.12197	0.089110	0.322490	0.449313	-0.05579		
		5	0.104111	0.396873	0.511790	-0.09900	0.052452	0.274609	0.354410	-0.04487		
		6	0.071720	0.352462	0.441284	-0.08382	0.036947	0.243803	0.305414	-0.03751		
		7	0.053410	0.320559	0.394999	-0.07321	0.028192	0.221690	0.273287	-0.03239		
		8	0.041644	0.296135	0.361267	-0.06534	0.022532	0.204769	0.249893	-0.02864		
		9	0.033502	0.276626	0.335151	-0.05921	0.018569	0.191258	0.231790	-0.02576		
		10	0.027592	0.260564	0.314105	-0.05425	0.015641	0.180137	0.217207	-0.02349		
		5	1	5	0.165311	0.185762	0.682101	0.038602	0.068527	0.128403	0.473039	0.007787
		6	0.128166	0.169597	0.547043	0.035389	0.051311	0.117219	0.378909	0.006896		
		7	0.110052	0.157029	0.475998	0.035306	0.042581	0.108526	0.329508	0.006820		
8	0.099086	0.146897	0.428730	0.036107	0.037070	0.101518	0.296682	0.006952				
9	0.091832	0.138503	0.393912	0.037274	0.033237	0.095714	0.272522	0.007153				
10	0.086817	0.131400	0.366721	0.038630	0.030414	0.090803	0.253665	0.007380				
2	5	5	0.173400	0.295213	0.682101	-0.02796	0.085162	0.204150	0.473039	-0.01141		
		6	0.119750	0.266589	0.547043	-0.02609	0.059439	0.184323	0.378909	-0.01040		
		7	0.092013	0.245001	0.475998	-0.02461	0.046480	0.169376	0.329508	-0.00933		
		8	0.073911	0.227956	0.428730	-0.02382	0.038296	0.157579	0.296682	-0.00846		

Table 2: Continued.

		$\tau = 2, \xi = 5, \delta = 0.5$					$\tau = 2, \xi = 10, \delta = 0.5$				
s	r	n	$\mu_{r,s;n}$	$E(X_{r;n})$	$E(X_{s;n})$	$\sigma_{r,s;n}$	$\mu_{r,s;n}$	$E(X_{r;n})$	$\mu_{r,s;n}$	$\sigma_{r,s;n}$	
3	1	9	0.060674	0.214051	0.393912	-0.02364	0.032550	0.147956	0.272522	-0.00777	
		10	0.050264	0.202423	0.366721	-0.02397	0.028244	0.139911	0.253665	-0.00725	
		5	0.202606	0.396873	0.682101	-0.06810	0.099405	0.274609	0.473039	-0.03050	
		6	0.133799	0.352462	0.547043	-0.05901	0.065762	0.243803	0.378909	-0.02662	
		7	0.100903	0.320559	0.475998	-0.05168	0.049799	0.221690	0.329508	-0.02325	
	4	8	0.080899	0.296135	0.428730	-0.04606	0.040135	0.204769	0.296682	-0.02062	
		9	0.067311	0.276626	0.393912	-0.04166	0.033574	0.191258	0.272522	-0.01855	
		10	0.057471	0.260564	0.366721	-0.03808	0.028798	0.180137	0.253665	-0.01690	
		5	0.197590	0.511790	0.682101	-0.15150	0.097898	0.354410	0.473039	-0.06975	
		6	0.115810	0.441284	0.547043	-0.12559	0.058351	0.305414	0.378909	-0.05737	
6	1	7	0.080302	0.394999	0.475998	-0.10772	0.041419	0.273287	0.329508	-0.04863	
		8	0.059934	0.361267	0.428730	-0.09495	0.031757	0.249893	0.296682	-0.04238	
		9	0.046659	0.335151	0.393912	-0.08536	0.025457	0.231790	0.272522	-0.03771	
		10	0.037334	0.314105	0.366721	-0.07786	0.021014	0.217207	0.253665	-0.03408	
		6	0.173540	0.169597	0.709112	0.053277	0.069056	0.117219	0.491865	0.011400	
	2	7	0.138937	0.157029	0.575461	0.048573	0.053325	0.108526	0.398669	0.010059	
		8	0.122004	0.146897	0.504359	0.047915	0.045185	0.101518	0.349204	0.009735	
		9	0.111775	0.138503	0.456584	0.048537	0.039958	0.095714	0.316010	0.009712	
		10	0.105094	0.131400	0.421103	0.049761	0.036274	0.090803	0.291379	0.009816	
		6	0.169539	0.266589	0.709112	-0.01950	0.083624	0.184323	0.491865	-0.00704	
7	1	7	0.121311	0.245001	0.575461	-0.01968	0.060716	0.169376	0.398669	-0.00681	
		8	0.095139	0.227956	0.504359	-0.01983	0.048693	0.157579	0.349204	-0.00633	
		9	0.077345	0.214051	0.456584	-0.02039	0.040853	0.147956	0.316010	-0.00590	
		10	0.063859	0.202423	0.421103	-0.02138	0.035208	0.139911	0.291379	-0.00556	
		6	0.206018	0.352462	0.709112	-0.04392	0.100332	0.243803	0.491865	-0.01959	
	2	7	0.145146	0.320559	0.575461	-0.03932	0.070483	0.221690	0.398669	-0.01790	
		8	0.114479	0.296135	0.504359	-0.03488	0.055437	0.204769	0.349204	-0.01607	
		9	0.095232	0.276626	0.456584	-0.03107	0.045952	0.191258	0.316010	-0.01449	
		10	0.081919	0.260564	0.421103	-0.02781	0.039327	0.180137	0.291379	-0.01316	
		6	0.223267	0.441284	0.709112	-0.08965	0.110002	0.305414	0.491865	-0.04022	
8	1	7	0.148164	0.394999	0.575461	-0.07914	0.073256	0.273287	0.398669	-0.03569	
		8	0.111855	0.361267	0.504359	-0.07035	0.055703	0.249893	0.349204	-0.03156	
		9	0.089461	0.335151	0.456584	-0.06356	0.045006	0.231790	0.316010	-0.02824	
		10	0.073972	0.314105	0.421103	-0.05830	0.037697	0.217207	0.291379	-0.02559	
		6	0.211616	0.547043	0.709112	-0.17630	0.104759	0.378909	0.491865	-0.08161	
	2	7	0.125475	0.475998	0.575461	-0.14844	0.063092	0.329508	0.398669	-0.06827	
		8	0.087690	0.428730	0.504359	-0.12854	0.045094	0.296682	0.349204	-0.05851	
		9	0.065827	0.393912	0.456584	-0.11403	0.034744	0.272522	0.316010	-0.05138	
		10	0.051487	0.366721	0.421103	-0.10294	0.027951	0.253665	0.291379	-0.04596	
		7	0.182405	0.157029	0.731387	0.067556	0.069687	0.108526	0.507398	0.014621	
9	1	8	0.149550	0.146897	0.599162	0.061535	0.055063	0.101518	0.415157	0.012917	
		9	0.133567	0.138503	0.528246	0.060403	0.047410	0.095714	0.365801	0.012398	
		10	0.124002	0.131400	0.480238	0.060899	0.042442	0.090803	0.332431	0.012256	
		7	0.163819	0.245001	0.731387	-0.01537	0.081635	0.169376	0.507398	-0.00431	
		8	0.119737	0.227956	0.599162	-0.01685	0.060918	0.157579	0.415157	-0.00450	
	2	9	0.094876	0.214051	0.528246	-0.01820	0.049741	0.147956	0.365801	-0.00438	
		10	0.077366	0.202423	0.480238	-0.01984	0.042276	0.139911	0.332431	-0.00423	
		7	0.205092	0.320559	0.731387	-0.02936	0.099057	0.221690	0.507398	-0.01343	
		8	0.150629	0.296135	0.599162	-0.02680	0.072334	0.204769	0.415157	-0.01268	
		9	0.122355	0.276626	0.528246	-0.02377	0.058356	0.191258	0.365801	-0.01161	
4	10	0.104276	0.260564	0.480238	-0.02086	0.049294	0.180137	0.332431	-0.01059		
	7	0.227234	0.320559	0.731387	-0.00722	0.111556	0.273287	0.507398	-0.02711		
	8	0.160109	0.296135	0.599162	-0.01732	0.078663	0.249893	0.415157	-0.02508		
	9	0.125850	0.276626	0.528246	-0.02028	0.062003	0.231790	0.365801	-0.02279		
10	0.103939	0.260564	0.480238	-0.02119	0.051445	0.217207	0.332431	-0.02076			

Table 2: Continued.

		$\tau = 2, \xi = 5, \delta = 0.5$					$\tau = 2, \xi = 10, \delta = 0.5$				
s	r	n	$\mu_{r,s;n}$	$E(X_{r;n})$	$E(X_{s;n})$	$\sigma_{r,s;n}$	$\mu_{r,s;n}$	$E(X_{r;n})$	$\mu_{r,s;n}$	$\sigma_{r,s;n}$	
8	5	7	0.240059	0.475998	0.731387	-0.10808	0.118411	0.329508	0.507398	-0.04878	
		8	0.160346	0.428730	0.599162	-0.09653	0.079330	0.296682	0.415157	-0.04384	
		9	0.121653	0.393912	0.528246	-0.08643	0.060576	0.272522	0.365801	-0.03911	
	6	10	0.097711	0.366721	0.480238	-0.07840	0.049095	0.253665	0.332431	-0.03523	
		7	0.223156	0.575461	0.731387	-0.19773	0.110362	0.398669	0.507398	-0.09192	
		8	0.133614	0.504359	0.599162	-0.16858	0.067043	0.349204	0.415157	-0.07793	
	1	9	0.094015	0.456584	0.528246	-0.14717	0.048208	0.316100	0.362580	-0.06739	
		10	0.070923	0.421103	0.480238	-0.13131	0.037307	0.291379	0.332431	-0.05956	
		8	0.192134	0.146897	0.750277	0.081920	0.070480	0.101518	0.520575	0.017632	
	2	9	0.160463	0.138503	0.619424	0.074671	0.056695	0.095714	0.429259	0.015609	
		10	0.145243	0.131400	0.548821	0.073128	0.049442	0.090803	0.380103	0.014928	
		8	0.156885	0.227956	0.750277	-0.01415	0.079554	0.157579	0.520575	-0.00248	
	3	9	0.116038	0.214051	0.619424	-0.01655	0.060572	0.147956	0.429259	-0.00294	
		10	0.092204	0.202423	0.548821	-0.01889	0.050124	0.139911	0.380103	-0.00306	
		8	0.203197	0.296135	0.750277	-0.01899	0.097193	0.204769	0.520575	-0.00940	
	4	9	0.153848	0.276626	0.619424	-0.01750	0.072961	0.191258	0.429259	-0.00914	
		10	0.127820	0.260564	0.548821	-0.01518	0.059979	0.180137	0.380103	-0.00849	
		8	0.225773	0.361267	0.750277	-0.04528	0.110449	0.249893	0.520575	-0.01964	
	5	9	0.165186	0.35151	0.619424	-0.04241	0.080817	0.231790	0.429259	-0.01868	
		10	0.133147	0.314105	0.548821	-0.03924	0.065255	0.217207	0.380103	-0.01731	
		8	0.245110	0.428730	0.750277	-0.07656	0.120610	0.296682	0.520575	-0.03384	
	6	9	0.173431	0.393912	0.619424	-0.07057	0.085377	0.272522	0.429259	-0.03161	
		10	0.136810	0.366721	0.548821	-0.06445	0.067482	0.253665	0.380103	-0.02894	
		8	0.253982	0.504359	0.750277	-0.12443	0.125347	0.349204	0.520575	-0.05644	
	7	9	0.170589	0.456584	0.619424	-0.11223	0.084412	0.316010	0.429259	-0.05124	
		10	0.129956	0.421103	0.548821	-0.10115	0.064702	0.291379	0.380103	-0.04605	
		8	0.232942	0.504359	0.750277	-0.14547	0.115084	0.415157	0.520575	-0.10104	
	8	9	0.140652	0.456584	0.619424	-0.14217	0.070427	0.365801	0.429259	-0.08660	
		10	0.099571	0.421103	0.548821	-0.13154	0.050909	0.332431	0.380103	-0.07545	
		9	0.202844	0.138503	0.766633	0.096663	0.071447	0.095714	0.531990	0.020528	
	9	10	0.171939	0.131400	0.637074	0.088228	0.058309	0.090803	0.441548	0.018215	
		9	0.148835	0.214051	0.766633	-0.01526	0.077490	0.147956	0.531990	-0.00122	
		10	0.110573	0.202423	0.637074	-0.01839	0.059909	0.139911	0.441548	-0.00187	
	10	9	0.201458	0.276626	0.766633	-0.01061	0.095239	0.191258	0.531990	-0.00651	
		10	0.156264	0.260564	0.637074	-0.00973	0.073008	0.180137	0.441548	-0.00653	
		9	0.222547	0.335151	0.766633	-0.03439	0.108528	0.231790	0.531990	-0.01478	
	1	10	0.167173	0.314105	0.637074	-0.03294	0.081537	0.217207	0.441548	-0.01437	
		9	0.244147	0.393912	0.766633	-0.05784	0.119779	0.272522	0.531990	-0.02520	
		10	0.179218	0.366721	0.637074	-0.05441	0.087879	0.253665	0.441548	-0.02413	
	2	9	0.260024	0.456584	0.766633	-0.09001	0.128154	0.316010	0.531990	-0.03996	
		10	0.184607	0.421103	0.637074	-0.08367	0.091033	0.291379	0.441548	-0.03762	
		9	0.265867	0.528246	0.766633	-0.13910	0.131234	0.365801	0.531990	-0.06337	
	3	10	0.179461	0.480238	0.637074	-0.12649	0.088775	0.332431	0.441548	-0.05801	
		9	0.241423	0.619424	0.766633	-0.23345	0.119156	0.429259	0.531990	-0.10921	
		10	0.146845	0.548821	0.637074	-0.20279	0.073382	0.380103	0.441548	-0.09445	
	4	9	0.214615	0.131400	0.781029	0.111988	0.072587	0.090803	0.542039	0.023369	
		10	0.139577	0.202423	0.781029	-0.01852	0.075465	0.139911	0.542039	-0.00037	
		9	0.200369	0.260564	0.781029	-0.00314	0.093369	0.180137	0.542039	-0.00427	
5	10	0.218699	0.314105	0.781029	-0.02663	0.106389	0.217207	0.542039	-0.01135		
	9	0.241199	0.366721	0.781029	-0.04522	0.117950	0.253665	0.542039	-0.01955		
	10	0.259491	0.421103	0.781029	-0.06940	0.127622	0.291379	0.542039	-0.03032		
6	9	0.272860	0.480238	0.781029	-0.10222	0.134607	0.332431	0.542039	-0.04558		
	10	0.276212	0.548821	0.781029	-0.15243	0.136336	0.380103	0.542039	-0.06969		
	9	0.248896	0.637074	0.781029	-0.24868	0.122731	0.441548	0.542039	-0.11661		

Table 3: Coefficients of the BLUEs of the location parameter for $\tau = 2, \delta = 0.5$.

ξ	n	c	$a_i, i = 1, 2, 3, \dots, (n - c)$							
5	4	0	0.60056	0.52054	-0.0249	-0.0962				
		1	1.14951	-0.0022	-0.1473					
		2	1.46179	-0.4618						
	8	0	0.94978	-0.0079	-0.0355	0.11756	0.04068	-0.0357	-0.0211	
			-0.0079							
		1	0.99500	-0.0542	-0.0168	0.13125	0.02120	-0.0520	-0.0245	
2		1.10008	-0.1932	0.02847	0.16721	-0.0256	-0.0769			
3		-0.1442	3.22795	-1.6960	-1.0836	0.69588				
4		1.22957	-5E-05	-0.2371	0.00762					
10	6	5	1.20093	0.03644	-0.2374					
		6	1.55875	-0.5587						
		0	0.84633	0.46665	-0.1943	0.02331	0.04968	-0.1587	-0.0944	
			0.02940	0.02612	0.00600					
		1	0.87109	0.40252	-0.1787	0.02622	0.03714	-0.1427	-0.0747	
			0.03345	0.02564						
	2	0.92580	0.23990	-0.1220	0.04079	0.00905	-0.0958	-0.0266		
		0.02881								
	3	1.14784	-0.0594	-0.0458	0.05472	-0.0716	-0.0521	0.02645		
	4	0.97357	0.11809	-0.0719	0.05459	-0.0127	-0.0617			
5	1.28460	-0.3778	0.12364	0.10493	-0.1354					
6	1.19440	0.06579	-0.2574	-0.0028						

Table 4: Coefficients of the BLUEs of the scale parameter for $\tau = 2, \delta = 0.5$.

ξ	n	c	$b_i, i = 1, 2, 3, \dots, (n - c)$						
5	4	0	-0.19072	0.15597	0.04744	-0.01269			
		1	-0.11835	0.08705	0.03130				
		2	-0.18472	0.18472					
	8	0	-0.11319	0.09364	0.00620	-0.02865	0.01231	0.02046	0.00619
			0.00305						
		1	-0.13061	0.11148	-0.00102	-0.03392	0.01981	0.02676	0.00750
2		-0.16282	0.15410	-0.01488	-0.04495	0.03417	0.03438		
3		0.39367	-1.37593	0.75633	0.51444	-0.28851			
4		-0.17592	-0.03759	0.15149	0.06202				
10	6	5	-0.40913	0.25959	0.14954				
		6	-0.63456	0.63456					
		0	-0.05656	-0.13237	0.10417	0.01245	-0.02274	0.06273	0.04769
			-0.00559	-0.00908	-0.00070				
		1	-0.05946	-0.12487	0.10235	0.01211	-0.02127	0.06086	0.04538
			-0.00606	-0.00902					
	2	-0.07871	-0.06765	0.08238	0.00698	-0.01139	0.04436	0.02846	
		-0.00443							
	3	-0.11285	-0.02163	0.07068	0.00484	0.00101	0.03764	0.02030	
	4	-0.24665	0.11466	0.05071	0.00474	0.04621	0.03032		
5	-0.39958	0.35850	-0.04542	-0.02001	0.10651				
6	-0.32861	0.00947	0.25437	0.06478					

Table 5: Variances and covariance of the BLUEs when $\tau = 0.5$, $\delta = 2$ and $\mu = 0$ and $\sigma = 1$.

ξ	n	c	$Var(\mu^*)$	$Var(\sigma^*)$	$Cov(\mu^*, \sigma^*)$
0.5	4	0	17.65257	1.410936	-0.566364
		1	14.02154	1.347831	-1.045046
		2	4.033713	0.896757	1.077511
	8	0	2.337863	0.846369	-0.520900
		1	2.167946	0.821161	-0.455453
		2	1.696970	0.776917	-0.311100
		3	20.75401	4.588457	-8.833814
		4	4.587047	1.809407	-2.130921
		5	4.594606	2.310717	-2.069361
10	0	2.750427	1.010938	-1.077970	
	1	2.628375	1.009270	-1.063700	
	2	2.206944	0.957081	-0.915396	
	3	1.727925	0.945757	-0.841744	
	4	1.848615	1.016895	-0.749085	
	5	0.544039	0.701471	-0.107607	
	6	3.260254	2.383097	-2.244815	
	7	3.261429	3.000961	-2.271760	
	8	0.140384	0.562046	0.487219	

5. Application with example

Here, we illustrate the applicability of the EPL distribution among a set of classical and recent models containing Pareto, generalized Lindley (GL) and type-II exponentiated log-logistic distribution (TIIELL), based on a set of goodness-of-fit statistics. We compare the goodness-of-fit of the models with the Akaike Information Criterion (AIC), Bayesian Information Criterion (BIC) goodness-of-fit statistics. Further, we get the Kolmogorov-Smirnov (K-S) statistics with their corresponding p-value. In general, the smaller the values of these statistics and the larger value of the p-value, the better the fit to the data.

Description of the data is as follows: The data set refers to a study on the vinyl chloride from clean upgradient monitoring wells in mg/L, which was studied by Bhaumik et al. (2009). The data are:

5.1 1.2 1.3 0.6 0.5 2.4 0.5 1.1 8.0 0.8 0.4 0.6 0.9 0.4 2.0 0.5 5.3
 3.2 2.7 2.9 2.5 2.3 1.0 0.2 0.1 0.1 1.8 0.9 2.0 4.0 6.8 1.2 0.4 0.2

In this data set, we select a random sample of size 10 and data are: 0.1, 1.1, 0.9, 2.3, 1.3, 2.5, 0.4, 2.0, 0.5, 3.2. By using the EPL distribution in Eq. (1) for the given sample, we have the maximum likelihood estimate of $\tau_{ML} = 1.2945$, $\delta_{ML} = 0.5514$ and $\xi_{ML} = 0.8099$. For this data set, a comparison of the EPL distribution with other classical and recent distributions is done as follows. We obtain the K-S, p-value, AIC and BIC statistics. The obtained results are reported in Table- 6. From this table, the smallest value of K-S, AIC and BIC and the largest value of p-value are obtained for the EPL distribution. Hence, we conclude

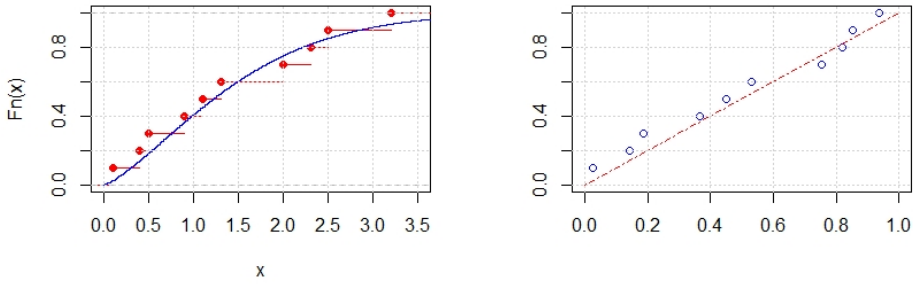


Figure 1: ECDF and Q-Q plot of the EPL distribution for the data set

that the EPL distribution provides the best fit among the compared distributions. Figure 1 shows Q-Q plot of the sample.

By using Tables 4 and 5, we have

$$\mu^* = \sum_{r=1}^n a_r X_{r:n} = 0.20081 \quad \text{and} \quad \sigma^* = \sum_{r=1}^n b_r X_{r:n} = 0.36501.$$

Table 6: The values of K-S, p- value, AIC and BIC statistics for some models fitted to the vinyl chloride data.

Distribution	K-S	p-value	AIC	BIC
Pareto	0.19982	0.1332	121.2704	122.7968
GL	0.11628	0.7474	116.2222	119.2747
THIELL	0.10401	0.8556	116.1526	119.2053
EPL	0.08219	0.9757	116.0253	120.6044

6. Conclusion

In this paper, we have obtained the explicit expression for single and double moments of order statistics. Also, these results are reduced for special cases of order statistics of the power Lindley distribution and the Lindley distribution, which is obtained by Kumar and Goyal (2019a) and Sultan and AL-Thubyani (2016) respectively. The single and double moments of order statistics can be used to obtain BLUEs of the location and scale parameters. A simulation study is conducted for calculating BLUEs for the location and scale parameters based on Type-II right-censored samples. Finally, one real data set has been used to obtain the MLEs of the model parameters, BLUEs of the parameters and the EPL distribution is also compared with some existing distributions. We conclude that the EPL distribution provides the best fit among the compared distributions.

Acknowledgments

The authors would like to express thanks to the Editor and anonymous referees for useful suggestions and comments which have improved the presentation of the manuscript.

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