



## Original Article



## Novel imputation methods under stratified simple random sampling

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## ABSTRACT

This paper addresses some classes of combined and separate imputation methods (CSIMs) of the population mean under stratified simple random sampling (SSRS) along with their characteristics. To the best of our knowledge, these imputation methods (IMs) have yet not been studied by any author under SSRS, hence these IMs are called 'novel'. In addition, the existing CSIMs are distinguished as the members of the suggested CSIMs, respectively. The theoretical conditions under which the proposed IMs perform better are obtained by comparing the proposed IMs with the existing IMs. To validate the theoretical findings, the numerical and simulation studies are conducted on real and artificial populations, respectively.

## 1. Introduction

In a sample survey, it is widely recognized that the complete information concerning any situation or episode is essential to make inferences. Incomplete information or missing values in the data set may impair the whole inference. The most familiar method which is used till date to tackle the issue of missing values is imputation. Many articles have been published so far to estimate the population mean in the presence of missing values utilizing simple random sampling (SRS). [1] mooted three key strategies of missing values, namely, missing at random (MAR), observed at random (OAR), and parameter distribution (PD). [2] marked a difference between missing at random and missing completely at random (MCAR) strategy. Subsequently, [3] suggested ratio category of estimators in case of missing values. In the presence of missing data, [4] developed an improved population mean imputation method. [5] suggested some optimal imputations for estimating population mean in case of missing data. [6] utilized moments of higher order of an auxiliary variable and suggested regression type IMs under SRS, whereas [7] considered higher order moments and suggested improved form of regression type IMs. [8] developed a generalized class of estimators utilizing the dual of the supplementary variable in the case of

non-response. [9] introduced composite imputation consisting of mean estimators utilizing robust quantile regression. [10] suggested logarithmic IMs using MCAR strategy. Recently, utilizing MCAR strategy, [11] and [12] suggested some IMs based on single and multi-auxiliary information under ranked set sampling. The basic concentration of this article is to discuss the MCAR strategy in the area of socio-economic investigations where sample units can be distinguished very cheaply.

It is well known that the stratified simple random sampling (SSRS) is better representative of a heterogeneous population than SRS. The SSRS enhances the efficiency of the estimators by separating the population into homogeneous strata over the sampling units. A new estimator for mean under SSRS was suggested by [13]. [14] considered SSRS to estimate the population mean using auxiliary characters. An effective exponential mean estimator for SSRS was studied by [15]. Utilizing SSRS, [16] developed an exponential ratio estimator of the median which is an alternative to the regression estimator of the median. [17] suggested dual utilization of auxiliary information for the estimation of finite population mean under SSRS, while [18] carried out a simulation study by utilizing dual auxiliary variable to estimate population mean under SSRS. In stratified two-phase sampling, [19] recommended using exponential ratio and product type estimators of the mean, whereas

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[20] suggested a generalized estimator for population mean utilizing auxiliary attribute. In SSRS, [21] investigated a few efficient types of estimators, however, [22] utilized bivariate auxiliary information and suggested a few improved types of estimators. [23] produced several enhanced classes of estimators by utilizing stratified ranked set sampling (SRSS), however, an efficient estimation of population mean under SRSS was presented by [24]. The issue of mean estimation under SRSS was discussed by [25]. A miniscule work has been done till date for estimating the population mean in case of missing data under SSRS. The estimate of the population mean under SSRS was taken into consideration by [26] using a few ratio type imputation approaches. Motivated by [27], [28] suggested a separate regression type estimator under SSRS, which is best linear unbiased estimator (BLUE). Under SSRS, there is no imputation method which compete with the BLUE. Apart from this, no study is available which considers CSIMs simultaneously. In the present study, these issues are taken into consideration, and the following objectives have been set:

- To propose novel CSIMs for the population mean under SSRS which compete with the BLUE.
- To compare theoretically the CSIMs with the corresponding conventional CSIMs.
- To perform an empirical study utilizing real and artificially rendered symmetric and skewed populations, respectively.

1.1. Methodology and notation

Consider a population  $\Theta = (\Theta_1, \Theta_2, \dots, \Theta_N)$  consisting of  $N$  units which is divided into  $L$  mutually exclusive and exhaustive strata with  $N_h$  units in the  $h^{th}$  stratum. Let a sample of size  $n_h$  units be chosen from the  $h^{th}$  stratum containing  $N_h$  elements to estimate the population mean. Let  $r$  be the number of units providing a response out of  $n$  selected units. The set of responding units is referred by  $R_u$  and the set of non-responding units is referred by  $\bar{R}_u$ . The notation of symbols employed throughout this study are defined below.

- $(N, n)$ ; population and sample size,
- $(N_h, n_h)$ ; population and sample size in stratum  $h$ ,
- $W_h = N_h/N$ ;  $h^{th}$  stratum's weight,
- $(\bar{y}_h = \sum_{i=1}^{n_h} y_{hi}/n_h, \bar{x}_h = \sum_{i=1}^{n_h} x_{hi}/n_h)$ ; sample mean of variables  $(y, x)$  in stratum  $h$ ,
- $(\bar{y}_{st} = \sum_{h=1}^L W_h y_h, \bar{x}_{st} = \sum_{h=1}^L W_h x_h)$ ; sample mean of variables  $(y, x)$ ,
- $(\bar{Y}_h = \sum_{i=1}^{N_h} y_{hi}/N_h, \bar{X}_h = \sum_{i=1}^{N_h} x_{hi}/N_h)$ ; population mean of variables  $(y, x)$  in stratum  $h$ ,
- $(\bar{Y} = \bar{Y}_{st} = \sum_{h=1}^L W_h Y_h, \bar{X} = \bar{X}_{st} = \sum_{h=1}^L W_h X_h)$ ; population mean of variables  $(y, x)$ ,
- $R = \bar{Y}/\bar{X}$ ; population ratio,
- $R_h = \bar{Y}_h/\bar{X}_h$ ;  $h^{th}$  stratum's population ratio,
- $(S_{y_h}^2 = (N_h - 1)^{-1} \sum_{i=1}^{N_h} (y_{hi} - \bar{Y}_h)^2, S_{x_h}^2 = (N_h - 1)^{-1} \sum_{i=1}^{N_h} (x_{hi} - \bar{X}_h)^2)$ ; population variance of variables  $(y, x)$  in stratum  $h$ ,
- $S_{xy_h} = (N_h - 1)^{-1} \sum_{i=1}^{N_h} (x_{hi} - \bar{X}_h)(y_{hi} - \bar{Y}_h)$ ; population covariance between variables  $x$  and  $y$  in stratum  $h$ ,
- $\rho_{xy_h} = S_{xy_h}/S_{x_h}S_{y_h}$ ; population correlation coefficient for the variables  $x$  and  $y$  in stratum  $h$ ,
- $(C_{y_h}, C_{x_h})$ ; population variation coefficients for the variables  $(y, x)$  in stratum  $h$ .

In this article, the undermentioned notation are considered to establish the attributes of the combined point estimators.

Let  $\bar{y}_{rst} = \bar{Y} + \epsilon_{0_{st}}, \bar{x}_{rst} = \bar{X} + \epsilon_{1_{st}}$  and  $\bar{x}_{n_{st}} = \bar{X} + \epsilon_{2_{st}}$  such that  $E(\epsilon_{0_{st}}) = E(\epsilon_{1_{st}}) = E(\epsilon_{2_{st}}) = 0, E(\epsilon_{0_{st}}^2) = \sum_{h=1}^L W_h^2 \gamma_h^* S_{y_h}^2 = I_0^*, E(\epsilon_{1_{st}}^2) = \sum_{h=1}^L W_h^2 \gamma_h^* S_{x_h}^2 = I_1^*, E(\epsilon_{2_{st}}^2) = E(\epsilon_{1_{st}}, \epsilon_{2_{st}}) = \sum_{h=1}^L W_h^2 \gamma_h \rho_{xy_h} S_{x_h} S_{y_h} = I_{01}^*$  and  $E(\epsilon_{0_{st}}, \epsilon_{2_{st}}) = \sum_{h=1}^L W_h^2 \gamma_h \rho_{xy_h} S_{x_h} S_{y_h} = I_{01}$ , where  $\gamma_h^* = (1/r_h) - (1/N_h)$  and  $\gamma_h = (1/n_h) - (1/N_h)$ .

The undermentioned notation are considered to determine the attributes of the separate point estimators.

Let  $\bar{y}_{rh} = \bar{Y}_h + e_{0_h}, \bar{x}_{rh} = \bar{X}_h + e_{1_h}, \bar{x}_{n_h} = \bar{X}_h + e_{2_h}$  such that  $E(e_{0_h}) = E(e_{1_h}) = E(e_{2_h}) = 0, E(e_{0_h}^2) = \gamma_h^* S_{y_h}^2 = J_0^*, E(e_{1_h}^2) = \gamma_h^* S_{x_h}^2 = J_1^*, E(e_{2_h}^2) = E(e_{1_h}, e_{2_h}) = \gamma_h S_{x_h}^2 = J_{01}, E(e_{0_h}, e_{1_h}) = \gamma_h^* \rho_{xy_h} S_{x_h} S_{y_h} = J_{01}^*$  and  $E(e_{0_h}, e_{2_h}) = \gamma_h \rho_{xy_h} S_{x_h} S_{y_h} = J_{01}$ .

The following sections make up the article's structure. We examine the commonly used CSIMs in Section 2. We propose some CSIMs in Section 3. Section 4 provides a theoretical comparison between existing and suggested IMs. Section 5 provides the simulation study along with the discussion of simulation results. The illustration of the proposed methods is shown in Section 6. Section 7 provides the conclusion of the study.

2. Existing imputation methods

The mean IM under SSRS is

$$y_{.im} = \begin{cases} y_i & \text{if } i \in R_u \\ \bar{y}_{rst} & \text{if } i \in \bar{R}_u \end{cases}$$

The sequent estimator is

$$T_m = \bar{y}_{rst}$$

where the stratified sample mean of study variable  $y$  is given by  $\bar{y}_{rst} =$

$$\sum_{h=1}^L W_h \bar{y}_h.$$

Further, we consider some prominent commonly used CSIMs.

2.1. Combined imputation methods

The IMs are divided into following strategies in the availability of auxiliary informations.

Strategy I: If  $\bar{X}$  is known and  $\bar{x}_{n_{st}}$  is used.

Strategy II: If  $\bar{X}$  is known and  $\bar{x}_{rst}$  is used.

Strategy III: If  $\bar{X}$  is unknown and  $\bar{x}_{n_{st}}$ , and  $\bar{x}_{rst}$  are used.

The conventional combined ratio IMs are prescribed under SSRS as

Strategy I

$$y_{.iR_1}^c = \begin{cases} y_i & \text{if } i \in R_u \\ \frac{1}{n-r} \left( n\bar{y}_{rst} \frac{\bar{X}}{\bar{x}_{n_{st}}} - r\bar{y}_{rst} \right) & \text{if } i \in \bar{R}_u \end{cases}$$

Strategy II

$$y_{.iR_2}^c = \begin{cases} y_i & \text{if } i \in R_u \\ \frac{1}{n-r} \left( n\bar{y}_{rst} \frac{\bar{X}}{\bar{x}_{rst}} - r\bar{y}_{rst} \right) & \text{if } i \in \bar{R}_u \end{cases}$$

Strategy III

$$y_{.iR_3}^c = \begin{cases} y_i & \text{if } i \in R_u \\ \frac{1}{n-r} \left( n\bar{y}_{rst} \frac{\bar{x}_{n_{st}}}{\bar{x}_{rst}} - r\bar{y}_{rst} \right) & \text{if } i \in \bar{R}_u \end{cases}$$

The sequent estimators are given as follows:

$$T_{R_1}^c = \bar{y}_{rst} \frac{\bar{X}}{\bar{x}_{n_{st}}}$$

$$T_{R_2}^c = \bar{y}_{rst} \frac{\bar{X}}{\bar{x}_{rst}}$$

$$T_{R_3}^c = \bar{y}_{rst} \frac{\bar{x}_{n_{st}}}{\bar{x}_{rst}}$$

where  $\bar{x}_{nst} = \sum_{h=1}^L W_h \bar{x}_h$  is the sample mean of the auxiliary variable  $x$  under SSRS.

Motivated by [4], we define the classical regression IM under SSRS as *Strategy I*

$$y_{iDP1}^c = \begin{cases} \bar{y}_i & \text{if } i \in R_u \\ \bar{y}_{rst} + \frac{n}{n-r} b_1 (\bar{X} - \bar{x}_{nst}) & \text{if } i \in \bar{R}_u \end{cases}$$

*Strategy II*

$$y_{iDP2}^c = \begin{cases} \bar{y}_i & \text{if } i \in R_u \\ \bar{y}_{rst} + \frac{n}{n-r} b_2 (\bar{X} - \bar{x}_{rst}) & \text{if } i \in \bar{R}_u \end{cases}$$

*Strategy III*

$$y_{iDP3}^c = \begin{cases} \bar{y}_i & \text{if } i \in R_u \\ \bar{y}_{rst} + \frac{n}{n-r} b_3 (\bar{x}_{nst} - \bar{x}_{rst}) & \text{if } i \in \bar{R}_u \end{cases}$$

The sequent estimators are given as follows:

$$T_{DP1}^c = \bar{y}_{rst} + b_1 (\bar{X} - \bar{x}_{nst})$$

$$T_{DP2}^c = \bar{y}_{rst} + b_2 (\bar{X} - \bar{x}_{rst})$$

$$T_{DP3}^c = \bar{y}_{rst} + b_3 (\bar{x}_{nst} - \bar{x}_{rst})$$

where  $b_j, j = 1, 2, 3$  are the regression coefficients for the respective strategies.

On the lines of [27], [26], and [28], we investigate the following combined ratio categories of IMs based on SSRS as

*Strategy I*

$$y_{iS1}^c = \begin{cases} y_i & \text{if } i \in R_u \\ \frac{n}{n-r} \left[ \bar{y}_{rst} \left( \frac{\bar{X}}{\bar{x}_{nst}} \right)^{\beta_1} - \bar{y}_{rst} \right] & \text{if } i \in \bar{R}_u \end{cases}$$

$$y_{iS4}^c = \begin{cases} y_i & \text{if } i \in R_u \\ \frac{n}{n-r} \left[ \bar{y}_{rst} \left( \frac{\bar{X}}{\beta_4 \bar{x}_{nst} + (1-\beta_4) \bar{X}} \right) - \bar{y}_{rst} \right] & \text{if } i \in \bar{R}_u \end{cases}$$

*Strategy II*

$$y_{iS2}^c = \begin{cases} y_i & \text{if } i \in R_u \\ \frac{n}{n-r} \left[ \bar{y}_{rst} \left( \frac{\bar{X}}{\bar{x}_{rst}} \right)^{\beta_2} - \bar{y}_{rst} \right] & \text{if } i \in \bar{R}_u \end{cases}$$

$$y_{iS5}^c = \begin{cases} y_i & \text{if } i \in R_u \\ \frac{n}{n-r} \left[ \bar{y}_{rst} \left( \frac{\bar{X}}{\beta_5 \bar{x}_{rst} + (1-\beta_5) \bar{X}} \right) - \bar{y}_{rst} \right] & \text{if } i \in \bar{R}_u \end{cases}$$

*Strategy III*

$$y_{iS3}^c = \begin{cases} y_i & \text{if } i \in R_u \\ \frac{n}{n-r} \left[ \bar{y}_{rst} \left( \frac{\bar{x}_{nst}}{\bar{x}_{rst}} \right)^{\beta_3} - \bar{y}_{rst} \right] & \text{if } i \in \bar{R}_u \end{cases}$$

$$y_{iS6}^c = \begin{cases} y_i & \text{if } i \in R_u \\ \frac{n}{n-r} \left[ \bar{y}_{rst} \left( \frac{\bar{X}}{\beta_6 \bar{x}_{rst} + (1-\beta_6) \bar{x}_{nst}} \right) - \bar{y}_{rst} \right] & \text{if } i \in \bar{R}_u \end{cases}$$

The sequent estimators are given by

$$T_{S1}^c = \bar{y}_{rst} \left( \frac{\bar{X}}{\bar{x}_{nst}} \right)^{\beta_1}$$

$$T_{S2}^c = \bar{y}_{rst} \left( \frac{\bar{X}}{\bar{x}_{rst}} \right)^{\beta_2}$$

$$T_{S3}^c = \bar{y}_{rst} \left( \frac{\bar{x}_{nst}}{\bar{x}_{rst}} \right)^{\beta_3}$$

$$T_{S4}^c = \bar{y}_{rst} \left( \frac{\bar{X}}{\beta_4 \bar{x}_{nst} + (1-\beta_4) \bar{X}} \right)$$

$$T_{S5}^c = \bar{y}_{rst} \left( \frac{\bar{X}}{\beta_5 \bar{x}_{rst} + (1-\beta_5) \bar{X}} \right)$$

$$T_{S6}^c = \bar{y}_{rst} \left( \frac{\bar{X}}{\beta_6 \bar{x}_{rst} + (1-\beta_6) \bar{x}_{nst}} \right)$$

where  $\beta_i; i = 1, 2, \dots, 6$  are properly chosen scalars.

In Appendix A, the mean square error (MSE) of the subsequent estimators derived from various IMs is provided.

### 2.2. Separate imputation methods

The separate methods of imputation are classified into following strategies in the accessibility of auxiliary informations.

*Strategy I:* If  $\bar{X}_h$  is known and  $\bar{x}_{nh}$  is used.

*Strategy II:* If  $\bar{X}_h$  is known and  $\bar{x}_{rh}$  is used.

*Strategy III:* If  $\bar{X}_h$  is unknown and  $\bar{x}_{nh}, \bar{x}_{rh}$  are used.

The classical separate ratio type methods of imputation under SSRS is defined as

*Strategy I*

$$y_{iR1}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \frac{1}{n-r} \left( n \bar{y}_{rh} \frac{\bar{X}_h}{\bar{x}_{nh}} - r \bar{y}_{rh} \right) & \text{if } i \in \bar{R}_{uh} \end{cases}$$

*Strategy II*

$$y_{iR2}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \frac{1}{n-r} \left( n \bar{y}_{rh} \frac{\bar{X}_h}{\bar{x}_{rh}} - r \bar{y}_{rh} \right) & \text{if } i \in \bar{R}_{uh} \end{cases}$$

*Strategy III*

$$y_{iR3}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \frac{1}{n-r} \left( n \bar{y}_{rh} \frac{\bar{x}_{nh}}{\bar{x}_{rh}} - r \bar{y}_{rh} \right) & \text{if } i \in \bar{R}_{uh} \end{cases}$$

The sequent estimators are given by

$$T_{R1}^s = \sum_{h=1}^L W_h \bar{y}_{rh} \frac{\bar{X}_h}{\bar{x}_{nh}}$$

$$T_{R2}^s = \sum_{h=1}^L W_h \bar{y}_{rh} \frac{\bar{X}_h}{\bar{x}_{rh}}$$

$$T_{R3}^s = \sum_{h=1}^L W_h \bar{y}_{rh} \frac{\bar{x}_{nh}}{\bar{x}_{rh}}$$

Following [4], we define the separate regression IMs under SSRS as

*Strategy I*

$$y_{iDP1}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \bar{y}_{rh} + \frac{n}{n-r} b_{1h} (\bar{X}_h - \bar{x}_{nh}) & \text{if } i \in \bar{R}_{uh} \end{cases}$$

*Strategy II*

$$y_{iDP2}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \bar{y}_{rh} + \frac{n}{n-r} b_{2h} (\bar{X}_h - \bar{x}_{rh}) & \text{if } i \in \bar{R}_{uh} \end{cases}$$

*Strategy III*

$$y_{iDP3}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \bar{y}_{rh} + \frac{n}{n-r} b_{3h} (\bar{x}_{nh} - \bar{x}_{rh}) & \text{if } i \in \bar{R}_{uh} \end{cases}$$

The sequent separate estimators are given by

$$T_{DP_1}^s = \sum_{h=1}^L W_h [\bar{y}_{r_h} + b_{1_h} (\bar{X}_h - \bar{x}_{n_h})]$$

$$T_{DP_2}^s = \sum_{h=1}^L W_h [\bar{y}_{r_h} + b_{2_h} (\bar{X}_h - \bar{x}_{r_h})]$$

$$T_{DP_3}^s = \sum_{h=1}^L W_h [\bar{y}_{r_h} + b_{3_h} (\bar{x}_{n_h} - \bar{x}_{r_h})]$$

On the lines of [27], we investigate the following separate ratio categories of IMs based on SSRS as

Strategy I

$$y_{.is_1}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \left[ \frac{n}{n-r} \left[ \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\bar{x}_{n_h}} \right)^{\beta_{1_h}} - \bar{y}_{r_h} \right] \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

$$y_{.is_4}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \left[ \frac{n}{n-r} \left[ \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\beta_{4_h} \bar{x}_{n_h} + (1-\beta_{4_h}) \bar{X}_h} \right) - \bar{y}_{r_h} \right] \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

Strategy II

$$y_{.is_2}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \left[ \frac{n}{n-r} \left[ \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\bar{x}_{r_h}} \right)^{\beta_{2_h}} - \bar{y}_{r_h} \right] \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

$$y_{.is_5}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \left[ \frac{n}{n-r} \left[ \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\beta_{5_h} \bar{x}_{r_h} + (1-\beta_{5_h}) \bar{X}_h} \right) - \bar{y}_{r_h} \right] \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

Strategy III

$$y_{.is_3}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \left[ \frac{n}{n-r} \left[ \bar{y}_{r_h} \left( \frac{\bar{x}_{n_h}}{\bar{x}_{r_h}} \right)^{\beta_{3_h}} - \bar{y}_{r_h} \right] \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

$$y_{.is_6}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \left[ \frac{n}{n-r} \left[ \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\beta_{6_h} \bar{x}_{r_h} + (1-\beta_{6_h}) \bar{x}_{n_h}} \right) - \bar{y}_{r_h} \right] \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

The sequent estimators are given by

$$T_{S_1}^s = \sum_{h=1}^L W_h \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\bar{x}_{n_h}} \right)^{\beta_{1_h}}$$

$$T_{S_2}^s = \sum_{h=1}^L W_h \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\bar{x}_{r_h}} \right)^{\beta_{2_h}}$$

$$T_{S_3}^s = \sum_{h=1}^L W_h \bar{y}_{r_h} \left( \frac{\bar{x}_{n_h}}{\bar{x}_{r_h}} \right)^{\beta_{3_h}}$$

$$T_{S_4}^s = \sum_{h=1}^L W_h \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\beta_{4_h} \bar{x}_{n_h} + (1-\beta_{4_h}) \bar{X}_h} \right)$$

$$T_{S_5}^s = \sum_{h=1}^L W_h \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\beta_{5_h} \bar{x}_{r_h} + (1-\beta_{5_h}) \bar{X}_h} \right)$$

$$T_{S_6}^s = \sum_{h=1}^L W_h \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\beta_{6_h} \bar{x}_{r_h} + (1-\beta_{6_h}) \bar{x}_{n_h}} \right)$$

where  $\beta_{i_h}$ ;  $i = 1, 2, \dots, 6$  are properly chosen scalars.

In Appendix B, the minimum MSE of the sequent estimators made up of several separate IMs is listed.

### 3. Suggested imputation methods

In the previous section, the IMs do not compete with the regression IM which is BLUE. The nub of this section is to suggest a few novel combined and separate population mean methods of imputation in case of missing data in SSRS which can compete with the BLUE. Motivated by [5], [29], [21], [10], and [30], we propose the following CSIMs based on SSRS.

#### 3.1. Combined imputation methods

We suggest nine novel combined methods of imputation in the strategies described in the former section as

Strategy I

$$y_{.iak_1} = \begin{cases} \alpha_1 y_i & \text{if } i \in R_u \\ \alpha_1 \bar{y}_{r_{st}} + \frac{n\theta_1}{n-r} (\bar{x}_{n_{st}} - \bar{X}) & \text{if } i \in \bar{R}_u \end{cases}$$

$$y_{.iak_4} = \begin{cases} y_i & \text{if } i \in R_u \\ \left[ \frac{1}{n-r} \left[ n\alpha_4 \bar{y}_{r_{st}} \left( \frac{\bar{X}}{\bar{x}_{n_{st}}} \right)^{\theta_4} - r\bar{y}_{r_{st}} \right] \right] & \text{if } i \in \bar{R}_u \end{cases}$$

$$y_{.iak_7} = \begin{cases} y_i & \text{if } i \in R_u \\ \left[ \frac{1}{n-r} \left[ n\alpha_7 \bar{y}_{r_{st}} \left( \frac{\bar{X}}{\bar{X} + \theta_7 (\bar{x}_{n_{st}} - \bar{X})} \right) - r\bar{y}_{r_{st}} \right] \right] & \text{if } i \in \bar{R}_u \end{cases}$$

Strategy II

$$y_{.iak_2} = \begin{cases} \alpha_2 y_i & \text{if } i \in R_u \\ \alpha_2 \bar{y}_{r_{st}} + \frac{n\theta_2}{n-r} (\bar{x}_{r_{st}} - \bar{X}) & \text{if } i \in \bar{R}_u \end{cases}$$

$$y_{.iak_5} = \begin{cases} y_i & \text{if } i \in R_u \\ \left[ \frac{1}{n-r} \left[ n\alpha_5 \bar{y}_{r_{st}} \left( \frac{\bar{X}}{\bar{x}_{r_{st}}} \right)^{\theta_5} - r\bar{y}_{r_{st}} \right] \right] & \text{if } i \in \bar{R}_u \end{cases}$$

$$y_{.iak_8} = \begin{cases} y_i & \text{if } i \in R_u \\ \left[ \frac{1}{n-r} \left[ n\alpha_8 \bar{y}_{r_{st}} \left( \frac{\bar{X}}{\bar{X} + \theta_8 (\bar{x}_{r_{st}} - \bar{X})} \right) - r\bar{y}_{r_{st}} \right] \right] & \text{if } i \in \bar{R}_u \end{cases}$$

Strategy III

$$y_{.iak_3} = \begin{cases} \alpha_3 y_i & \text{if } i \in R_u \\ \alpha_3 \bar{y}_{r_{st}} + \frac{n\theta_3}{n-r} (\bar{x}_{r_{st}} - \bar{x}_{n_{st}}) & \text{if } i \in \bar{R}_u \end{cases}$$

$$y_{.iak_6} = \begin{cases} y_i & \text{if } i \in R_u \\ \left[ \frac{1}{n-r} \left[ n\alpha_6 \bar{y}_{r_{st}} \left( \frac{\bar{x}_{n_{st}}}{\bar{x}_{r_{st}}} \right)^{\theta_6} - r\bar{y}_{r_{st}} \right] \right] & \text{if } i \in \bar{R}_u \end{cases}$$

$$y_{.iak_9} = \begin{cases} y_i & \text{if } i \in R_u \\ \left[ \frac{1}{n-r} \left[ n\alpha_9 \bar{y}_{r_{st}} \left( \frac{\bar{x}_{n_{st}}}{\bar{x}_{n_{st}} + \theta_9 (\bar{x}_{n_{st}} - \bar{x}_{r_{st}})} \right) - r\bar{y}_{r_{st}} \right] \right] & \text{if } i \in \bar{R}_u \end{cases}$$

Under the above strategies, the point estimators are given as follows:

$$T_{ak_1}^c = \alpha_1 \bar{y}_{r_{st}} + \theta_1 (\bar{x}_{n_{st}} - \bar{X})$$

$$T_{ak_2}^c = \alpha_2 \bar{y}_{r_{st}} + \theta_2 (\bar{x}_{r_{st}} - \bar{X})$$

$$T_{ak_3}^c = \alpha_3 \bar{y}_{r_{st}} + \theta_3 (\bar{x}_{r_{st}} - \bar{x}_{n_{st}})$$

$$T_{ak_4}^c = \alpha_4 \bar{y}_{r_{st}} \left( \frac{\bar{X}}{\bar{x}_{n_{st}}} \right)^{\theta_4}$$

$$T_{ak_5}^c = \alpha_5 \bar{y}_{r_{st}} \left( \frac{\bar{X}}{\bar{x}_{r_{st}}} \right)^{\theta_5}$$

$$T_{ak_6}^c = \alpha_6 \bar{y}_{r_{st}} \left( \frac{\bar{x}_{n_{st}}}{\bar{x}_{r_{st}}} \right)^{\theta_6}$$

$$T_{ak_7}^c = \alpha_7 \bar{y}_{r_{st}} \left[ \frac{\bar{X}}{\bar{X} + \theta_7(\bar{x}_{n_{st}} - \bar{X})} \right]$$

$$T_{ak_8}^c = \alpha_8 \bar{y}_{r_{st}} \left[ \frac{\bar{X}}{\bar{X} + \theta_8(\bar{x}_{r_{st}} - \bar{X})} \right]$$

$$T_{ak_9}^c = \alpha_9 \bar{y}_{r_{st}} \left[ \frac{\bar{x}_{n_{st}}}{\bar{x}_{n_{st}} + \theta_9(\bar{x}_{r_{st}} - \bar{x}_{n_{st}})} \right]$$

where  $\alpha_j$  and  $\theta_j$ ;  $j = 1, 2, \dots, 9$  have been properly chosen scalars. Notably, the suggested combined IMs  $y_{i,ak_j}^c$ ,  $j = 1, 2, \dots, 9$  reduce to the current combined IMs for known values of scalars as

1. Conventional mean IM  $y_{i,m}^c$  for  $(\alpha_j, \theta_j; j = 4, 5, 6) = (1, 0)$ .
2. Classical ratio IM  $y_{i,r_j}^c$ ,  $j = 1, 2, 3$  for  $(\alpha_j, \theta_j; j = 4, 5, 6) = (1, 1)$ .
3. Imputation techniques envisaged on the lines of [27]  $y_{i,S_j}^c$ ,  $j = 1, 2, \dots, 6$  for  $(\alpha_j, \theta_j; j = 4, 5, \dots, 9) = (\alpha_j, 1)$ .
4. Imputation techniques envisaged on the lines of [4]  $y_{i,DP_j}^c$ ,  $j = 1, 2, 3$  for  $(\alpha_j, \theta_j; j = 1, 2, 3) = (1, b_j)$ .

### 3.1.1. MSE and minimum MSE of the proposed combined estimators

The MSE of the sequent estimators consisting of the suggested methods imputation is provided as

$$MSE(T_{ak_1}^c) = (\alpha_1 - 1)^2 \bar{Y}^2 + \alpha_1^2 \bar{Y}^2 I_0^* + \theta_1^2 \bar{X}^2 I_1 + 2\alpha_1 \theta_1 \bar{X} \bar{Y} I_{01}$$

$$MSE(T_{ak_2}^c) = (\alpha_2 - 1)^2 \bar{Y}^2 + \alpha_2^2 \bar{Y}^2 I_0^* + \theta_2^2 \bar{X}^2 I_1^* + 2\alpha_2 \theta_2 \bar{X} \bar{Y} I_{01}^*$$

$$MSE(T_{ak_3}^c) = \left[ (\alpha_3 - 1)^2 \bar{Y}^2 + \alpha_3^2 \bar{Y}^2 I_0^* + \theta_3^2 \bar{X}^2 \{I_1^* - I_1\} \right] + 2\alpha_3 \theta_3 \bar{X} \bar{Y} \{I_{01}^* - I_{01}\}$$

$$MSE(T_{ak_4}^c) = \bar{Y}^2 \left[ \frac{1 + \alpha_4^2 \{1 + I_0^* + \theta_4(2\theta_4 + 1)I_1 - 4\theta_4 I_{01}\}}{-2\alpha_4 \{1 - \theta_4 I_{01} + \frac{\theta_4(\theta_4 + 1)}{2} I_1\}} \right]$$

$$MSE(T_{ak_5}^c) = \bar{Y}^2 \left[ \frac{1 + \alpha_5^2 \{1 + I_0^* + \theta_5(2\theta_5 + 1)I_1^* - 4\theta_5 I_{01}^*\}}{-2\alpha_5 \{1 - \theta_5 I_{01}^* + \frac{\theta_5(\theta_5 + 1)}{2} I_1^*\}} \right]$$

$$MSE(T_{ak_6}^c) = \bar{Y}^2 \left[ \frac{1 + \alpha_6^2 \{1 + I_0^* + \theta_6(2\theta_6 + 1)(I_1^* - I_1) - 4\theta_6(I_{01}^* - I_{01})\}}{-2\alpha_6 \{1 - \theta_6(I_{01}^* - I_{01}) + \frac{\theta_6(\theta_6 + 1)}{2}(I_1^* - I_1)\}} \right]$$

$$MSE(T_{ak_7}^c) = \bar{Y}^2 \left[ \frac{1 + \alpha_7^2 \{1 + I_0^* + 3\theta_7^2 I_1 - 4\theta_7 I_{01}\}}{-2\alpha_7 \{1 + \theta_7^2 I_1 - \theta_7 I_{01}\}} \right]$$

$$MSE(T_{ak_8}^c) = \bar{Y}^2 \left[ \frac{1 + \alpha_8^2 \{1 + I_0^* + 3\theta_8^2 I_1^* - 4\theta_8 I_{01}^*\}}{-2\alpha_8 \{1 + \theta_8^2 I_1^* - \theta_8 I_{01}^*\}} \right]$$

$$MSE(T_{ak_9}^c) = \bar{Y}^2 \left[ \frac{1 + \alpha_9^2 \{1 + I_0^* + 3\theta_9^2(I_1^* - I_1) - 4\theta_9(I_{01}^* - I_{01})\}}{-2\alpha_9 \{1 + \theta_9^2(I_1^* - I_1) - \theta_9(I_{01}^* - I_{01})\}} \right]$$

The minimum MSE of the sequent estimators consisting of the suggested IMs is provided by

$$\min MSE(T_{ak_j}^c) = \bar{Y}^2(1 - \alpha_{j(opt)}) = \bar{Y}^2 \left( 1 - \frac{A_j^2}{B_j} \right); j = 1, 2, 3 \quad (3.1)$$

$$\min MSE(T_{ak_j}^c) = \bar{Y}^2 \left( 1 - \frac{A_j^2}{B_j} \right); j = 4, 5, 6, 7, 8, 9 \quad (3.2)$$

Appendix C contains the derivations of these MSE expressions along with brief annotations.

### 3.2. Separate imputation methods

We suggest the following new separate IMs under the strategies described primarily as Strategy I

$$y_{i,ak_1}^s = \begin{cases} \alpha_{1h} y_i & \text{if } i \in R_{uh} \\ \alpha_{1h} \bar{y}_{r_h} + \frac{n\theta_{1h}}{n-r} (\bar{x}_{n_h} - \bar{X}_h) & \text{if } i \in \bar{R}_{uh} \end{cases}$$

$$y_{i,ak_4}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \frac{1}{n-r} \left[ n\alpha_{4h} \bar{y}_{r_h} \left( \frac{\bar{x}_h}{\bar{x}_{n_h}} \right)^{\theta_{4h}} - r\bar{y}_{r_h} \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

$$y_{i,ak_7}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \frac{1}{n-r} \left[ n\alpha_{7h} \bar{y}_{r_h} \left( \frac{\bar{x}_h}{\bar{x}_h + \theta_{7h}(\bar{x}_{n_h} - \bar{X}_h)} \right) - r\bar{y}_{r_h} \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

#### Strategy II

$$y_{i,ak_2}^s = \begin{cases} \alpha_{2h} y_i & \text{if } i \in R_{uh} \\ \alpha_{2h} \bar{y}_{r_h} + \frac{n\theta_{2h}}{n-r} (\bar{x}_{r_h} - \bar{X}_h) & \text{if } i \in \bar{R}_{uh} \end{cases}$$

$$y_{i,ak_5}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \frac{1}{n-r} \left[ n\alpha_{5h} \bar{y}_{r_h} \left( \frac{\bar{x}_h}{\bar{x}_{r_h}} \right)^{\theta_{5h}} - r\bar{y}_{r_h} \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

$$y_{i,ak_8}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \frac{1}{n-r} \left[ n\alpha_{8h} \bar{y}_{r_h} \left( \frac{\bar{x}_h}{\bar{x}_h + \theta_{8h}(\bar{x}_{r_h} - \bar{X}_h)} \right) - r\bar{y}_{r_h} \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

#### Strategy III

$$y_{i,ak_3}^s = \begin{cases} \alpha_{3h} y_i & \text{if } i \in R_{uh} \\ \alpha_{3h} \bar{y}_{r_h} + \frac{n\theta_{3h}}{n-r} (\bar{x}_{r_h} - \bar{x}_{n_h}) & \text{if } i \in \bar{R}_{uh} \end{cases}$$

$$y_{i,ak_6}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \frac{1}{n-r} \left[ n\alpha_{6h} \bar{y}_{r_h} \left( \frac{\bar{x}_{n_h}}{\bar{x}_{r_h}} \right)^{\theta_{6h}} - r\bar{y}_{r_h} \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

$$y_{i,ak_9}^s = \begin{cases} y_i & \text{if } i \in R_{uh} \\ \frac{1}{n-r} \left[ n\alpha_{9h} \bar{y}_{r_h} \left( \frac{\bar{x}_{n_h}}{\bar{x}_{n_h} + \theta_{9h}(\bar{x}_{n_h} - \bar{x}_{r_h})} \right) - r\bar{y}_{r_h} \right] & \text{if } i \in \bar{R}_{uh} \end{cases}$$

The sequent estimators are given by

$$T_{ak_1}^s = \sum_{h=1}^L W_h [\alpha_{1h} \bar{y}_{r_h} + \theta_{1h} (\bar{x}_{n_h} - \bar{X}_h)]$$

$$T_{ak_2}^s = \sum_{h=1}^L W_h [\alpha_{2h} \bar{y}_{r_h} + \theta_{2h} (\bar{x}_{r_h} - \bar{X}_h)]$$

$$T_{ak_3}^s = \sum_{h=1}^L W_h [\alpha_{3h} \bar{y}_{r_h} + \theta_{3h} (\bar{x}_{r_h} - \bar{x}_{n_h})]$$

$$T_{ak_4}^s = \sum_{h=1}^L W_h \alpha_{4h} \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\bar{x}_{n_h}} \right)^{\theta_{4h}}$$

$$T_{ak_5}^s = \sum_{h=1}^L W_h \alpha_{5h} \bar{y}_{r_h} \left( \frac{\bar{X}_h}{\bar{x}_{r_h}} \right)^{\theta_{5h}}$$

$$T_{ak_6}^s = \sum_{h=1}^L W_h \alpha_{6h} \bar{y}_{r_h} \left( \frac{\bar{x}_{n_h}}{\bar{x}_{r_h}} \right)^{\theta_{6h}}$$

$$T_{ak_7}^s = \sum_{h=1}^L W_h \alpha_{7h} \bar{y}_{r_h} \left[ \frac{\bar{X}_h}{\bar{X}_h + \theta_{7h} (\bar{x}_{n_h} - \bar{X}_h)} \right]$$

$$T_{ak_8}^s = \sum_{h=1}^L W_h \alpha_{8h} \bar{y}_{r_h} \left[ \frac{\bar{X}_h}{\bar{X}_h + \theta_{8h} (\bar{x}_{r_h} - \bar{X}_h)} \right]$$

$$T_{ak_9}^s = \sum_{h=1}^L W_h \alpha_{9h} \bar{y}_{r_h} \left[ \frac{\bar{x}_{n_h}}{\bar{x}_{n_h} + \theta_{9h} (\bar{x}_{r_h} - \bar{x}_{n_h})} \right]$$

where  $\alpha_{jh}$  and  $\theta_{jh}$ ;  $j = 1, 2, \dots, 9$  are the properly chosen scalars. It is important to note that the suggested separate IMs  $y_{iakj}^s$ ,  $j = 1, 2, \dots, 9$  reduce to the existing separate IMs for known values of scalars as

1. Conventional mean IM  $y_{jm}^c$  for  $(\alpha_{jh}, \theta_{jh}; j = 4, 5, 6) = (1, 0)$ .
2. Classical ratio IM  $y_{iRjh}^c$ ,  $j = 1, 2, 3$  for  $(\alpha_{jh}, \theta_{jh}; j = 4, 5, 6) = (1, 1)$ .
3. Imputation techniques envisaged on the lines of [27]  $y_{iSjh}^c$ ,  $j = 1, 2, \dots, 6$  for  $(\alpha_{jh}, \theta_{jh}; j = 4, 5, \dots, 9) = (\alpha_{jh}, 1)$ .
4. Imputation techniques envisaged on the lines of [4]  $y_{iDPjh}^c$ ,  $j = 1, 2, 3$  for  $(\alpha_{jh}, \theta_{jh}; j = 1, 2, 3) = (1, b_{jh})$ .

### 3.2.1. MSE and minimum MSE of the proposed separate estimators

The MSE of the sequent estimators based on the suggested IMs is given by

$$\begin{aligned}
 &MSE(T_{ak_1}^s) \\
 &= \sum_{h=1}^L W_h^2 \left\{ (\alpha_{1h} - 1)^2 \bar{Y}_h^2 + \alpha_{1h}^2 \bar{Y}_h^2 J_0^* + \theta_{1h}^2 \bar{X}_h^2 J_1 + 2\alpha_{1h} \theta_{1h} \bar{X}_h \bar{Y}_h J_{01} \right\} \\
 &MSE(T_{ak_2}^s) \\
 &= \sum_{h=1}^L W_h^2 \left\{ (\alpha_{2h} - 1)^2 \bar{Y}_h^2 + \alpha_{2h}^2 \bar{Y}_h^2 J_0^* + \theta_{2h}^2 \bar{X}_h^2 J_1^* + 2\alpha_{2h} \theta_{2h} \bar{X}_h \bar{Y}_h J_{01}^* \right\} \\
 &MSE(T_{ak_3}^s) = \sum_{h=1}^L W_h^2 \left[ (\alpha_{3h} - 1)^2 \bar{Y}_h^2 + \alpha_{3h}^2 \bar{Y}_h^2 J_0^* + \theta_{3h}^2 \bar{X}_h^2 \{J_1^* - J_1\} \right. \\
 &\quad \left. + 2\alpha_{3h} \theta_{3h} \bar{X}_h \bar{Y}_h \{J_{01}^* - J_{01}\} \right] \\
 &MSE(T_{ak_4}^s) \\
 &= \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ 1 + \alpha_{4h}^2 \left\{ 1 + J_0^* + \theta_{4h} (2\theta_{4h} + 1) J_1 - 4\theta_{4h} J_{01} \right\} \right. \\
 &\quad \left. - 2\alpha_{4h} \left\{ 1 - \theta_{4h} J_{01} + \frac{\theta_{4h}(\theta_{4h} + 1)}{2} J_1 \right\} \right] \\
 &MSE(T_{ak_5}^s) \\
 &= \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ 1 + \alpha_{5h}^2 \left\{ 1 + J_0^* + \theta_{5h} (2\theta_{5h} + 1) J_1^* - 4\theta_{5h} J_{01}^* \right\} \right. \\
 &\quad \left. - 2\alpha_{5h} \left\{ 1 - \theta_{5h} J_{01}^* + \frac{\theta_{5h}(\theta_{5h} + 1)}{2} J_1^* \right\} \right] \\
 &MSE(T_{ak_6}^s) \\
 &= \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ 1 + \alpha_{6h}^2 \left\{ 1 + J_0^* + \theta_{6h} (2\theta_{6h} + 1) (J_1^* - J_1) \right\} \right. \\
 &\quad \left. - 4\theta_{6h} (J_{01}^* - J_{01}) - 2\alpha_{6h} \left\{ 1 - \theta_{6h} (J_{01}^* - J_{01}) + \frac{\theta_{6h}(\theta_{6h} + 1)}{2} (J_1^* - J_1) \right\} \right] \\
 &MSE(T_{ak_7}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ 1 + \alpha_{7h}^2 \left\{ 1 + J_0^* + 3\theta_{7h}^2 J_1 - 4\theta_{7h} J_{01} \right\} \right. \\
 &\quad \left. - 2\alpha_{7h} \left\{ 1 + \theta_{7h}^2 J_1 - \theta_{7h} J_{01} \right\} \right] \\
 &MSE(T_{ak_8}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ 1 + \alpha_{8h}^2 \left\{ 1 + J_0^* + 3\theta_{8h}^2 J_1^* - 4\theta_{8h} J_{01}^* \right\} \right. \\
 &\quad \left. - 2\alpha_{8h} \left\{ 1 + \theta_{8h}^2 J_1^* - \theta_{8h} J_{01}^* \right\} \right] \\
 &MSE(T_{ak_9}^s) \\
 &= \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ 1 + \alpha_{9h}^2 \left\{ 1 + J_0^* + 3\theta_{9h}^2 (J_1^* - J_1) \right\} \right. \\
 &\quad \left. - 4\theta_{9h} (J_{01}^* - J_{01}) - 2\alpha_{9h} \left\{ 1 + \theta_{9h}^2 (J_1^* - J_1) - \theta_{9h} (J_{01}^* - J_{01}) \right\} \right]
 \end{aligned}$$

The minimum MSEs of the sequent estimators consisting of the prof-ferred IMs is given as

$$\min MSE(T_{ak_j}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 (1 - \alpha_{jh(opt)}); \quad j = 1, 2, 3 \tag{3.3}$$

$$\min MSE(T_{ak_j}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left( 1 - \frac{A_{jh}^2}{B_{jh}} \right); \quad j = 4, 5, 6, 7, 8, 9 \tag{3.4}$$

Appendix C contains the derivations of these MSE expressions along with brief annotations.

## 4. Theoretical conditions

### 4.1. Combined imputation methods

The following theoretical conditions are obtained by comparing the minimum mean square error of the suggested combined IMs  $y_{iakj}^c$ ,  $j = 1, 2, \dots, 9$  provided in (3.1) and (3.2) with the minimum mean square error of the conventional combined IMs provided from (A.1) to (A.10).

$$\begin{aligned}
 &MSE(T_m) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^*, \quad j = 1, 2, \dots, 9 \\
 &MSE(T_{R_1}^c) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^* - I_1 + 2I_{01} \\
 &MSE(T_{R_2}^c) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^* - I_1^* + 2I_{01}^* \\
 &MSE(T_{R_3}^c) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^* + I_1 - I_1^* + 2(I_{01}^* - I_{01}) \\
 &MSE(T_{DP_1}^c) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^* + \frac{I_{01}^2}{I_1} \\
 &MSE(T_{DP_2}^c) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^* + \frac{I_{01}^*^2}{I_1^*} \\
 &MSE(T_{DP_3}^c) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^* + \frac{(I_{01}^* - I_{01})^2}{(I_1^* - I_1)} \\
 &MSE(T_{S_1}^c) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^* + \frac{I_{01}^2}{I_1} \\
 &MSE(T_{S_2}^c) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^* + \frac{I_{01}^*^2}{I_1^*} \\
 &MSE(T_{S_3}^c) > MSE(T_{ak_j}^c) \implies \frac{A_j^2}{B_j} > 1 - I_0^* + \frac{(I_{01}^* - I_{01})^2}{(I_1^* - I_1)}
 \end{aligned}$$

The proposed combined IMs will be reasonable provided these conditions hold.

### 4.2. Separate imputation methods

The following theoretical conditions are obtained by comparing the minimum mean square error of the suggested combined IMs  $y_{iakj}^s$ ,  $j = 1, 2, \dots, 9$  provided in (3.3) and (3.4) with the minimum mean square error of the conventional combined IMs provided in (A.1) to (B.19).

$$\begin{aligned}
 &MSE(T_m) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left( 1 - \frac{A_{jh}^2}{B_{jh}} \right) \\
 &\quad < \bar{Y}^2 I_0^*, \quad j = 1, 2, \dots, 9 \\
 &MSE(T_{R_1}^s) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left( 1 - \frac{A_{jh}^2}{B_{jh}} \right) \\
 &\quad < \sum_{h=1}^L W_h^2 \bar{Y}_h^2 [J_0^* + J_1 - 2J_{01}] \\
 &MSE(T_{R_2}^s) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left( 1 - \frac{A_{jh}^2}{B_{jh}} \right) \\
 &\quad < \sum_{h=1}^L W_h^2 \bar{Y}_h^2 [J_0^* + J_1^* - 2J_{01}^*]
 \end{aligned}$$

$$\begin{aligned}
 &MSE(T_{R_3}^s) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{jh}^2}{B_{jh}}\right) \\
 &< \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ \frac{J_0^* + J_1^* - J_1}{-2(J_{01}^* - J_{01})} \right] \\
 &MSE(T_{DP_1}^s) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{jh}^2}{B_{jh}}\right) \\
 &< \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{J_{01}^2}{J_1} \right] \\
 &MSE(T_{DP_2}^s) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{jh}^2}{B_{jh}}\right) \\
 &< \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{J_{01}^{*2}}{J_1^*} \right] \\
 &MSE(T_{DP_3}^s) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{jh}^2}{B_{jh}}\right) \\
 &< \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{(J_{01}^* - J_{01})^2}{(J_1^* - J_1)} \right] \\
 &MSE(T_{S_1}^s) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{jh}^2}{B_{jh}}\right) \\
 &< \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{J_{01}^2}{J_1} \right] \\
 &MSE(T_{S_2}^s) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{jh}^2}{B_{jh}}\right) \\
 &< \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{J_{01}^{*2}}{J_1^*} \right] \\
 &MSE(T_{S_3}^s) > MSE(T_{ak_j}^s) \implies \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left(1 - \frac{A_{jh}^2}{B_{jh}}\right) \\
 &< \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{(J_{01}^* - J_{01})^2}{(J_1^* - J_1)} \right]
 \end{aligned}$$

The suggested separate IMs can be reasonable provided the above conditions hold.

### 4.3. Comparison of suggested CSIMs

The following condition is obtained by comparing the minimum MSE of the suggested combined and separate classes of estimators provided in (3.1), (3.2) and (3.3), (3.4).

$$\begin{aligned}
 &minMSE(T_{ak_j}^c) - minMSE(T_{ak_j}^s) \\
 &= \sum_{h=1}^L \left[ (\bar{Y}^2 - W_h^2 \bar{Y}_h^2) - \left( \bar{Y}^2 \frac{A_j^2}{B_j} - W_h^2 \bar{Y}_h^2 \frac{A_{jh}^2}{B_{jh}} \right) \right] \quad (4.1)
 \end{aligned}$$

The last term of (4.1) is generally insignificant and it becomes faded in if:

1. the sequent estimators are conclusive;
2. within each stratum, if both the variables (x, y) are linearly related and the regression line goes through the origin.

The suggested combined estimators are to be strongly preferred having very small sample in every stratum, as the separate estimators outperform in every stratum with the exception of  $R_h$  becoming stable from

**Table 1**  
PRE of suggested combined estimators.

| Correlation coefficient            | 0.6      | 0.7      | 0.8      | 0.9      |
|------------------------------------|----------|----------|----------|----------|
| $T_m$                              | 100      | 100      | 100      | 100      |
| $x^* \sim N(2, 7)$                 |          |          |          |          |
| $y^* \sim N(9, 11)$                |          |          |          |          |
| <b>Strategy I</b>                  |          |          |          |          |
| $T_{R_1}^c$                        | 55.2264  | 55.6970  | 56.8299  | 58.5490  |
| $T_{DP_1}^c / T_{S_1}^c, i = 1, 4$ | 100.6625 | 100.8165 | 100.6375 | 100.5616 |
| $T_{ak_j}^c, j = 1, 7$             | 110.7267 | 110.8560 | 111.3526 | 112.1161 |
| $T_{ak_s}^c$                       | 111.4849 | 111.6909 | 112.1184 | 112.8627 |
| <b>Strategy II</b>                 |          |          |          |          |
| $T_{R_2}^c$                        | 36.8794  | 37.3240  | 38.4072  | 40.0865  |
| $T_{DP_2}^c / T_{S_2}^c, i = 1, 4$ | 101.4089 | 101.7394 | 101.3554 | 101.1931 |
| $T_{ak_j}^c, j = 2, 8$             | 111.4732 | 111.7789 | 112.0705 | 112.7476 |
| $T_{ak_s}^c$                       | 113.1226 | 113.6026 | 113.7344 | 114.3652 |
| <b>Strategy III</b>                |          |          |          |          |
| $T_{R_3}^c$                        | 52.6092  | 53.0839  | 54.2288  | 55.9712  |
| $T_{DP_3}^c / T_{S_3}^c, i = 3, 6$ | 100.7366 | 100.9080 | 100.7088 | 100.6244 |
| $T_{ak_j}^c, j = 3, 9$             | 110.8009 | 110.9475 | 111.4239 | 112.1789 |
| $T_{ak_s}^c$                       | 111.6458 | 111.8783 | 112.2773 | 113.0106 |
| $x^* \sim Exp(1.0)$                |          |          |          |          |
| $y^* \sim Exp(0.5)$                |          |          |          |          |
| <b>Strategy I</b>                  |          |          |          |          |
| $T_{R_1}^c$                        | 52.6742  | 49.0842  | 51.2637  | 55.3784  |
| $T_{DP_1}^c / T_{S_1}^c, i = 1, 4$ | 100.3880 | 100.4866 | 100.4600 | 100.5303 |
| $T_{ak_j}^c, j = 1, 7$             | 100.7498 | 100.7982 | 100.7993 | 100.9215 |
| $T_{ak_s}^c$                       | 100.7715 | 100.8209 | 100.8222 | 100.9477 |
| <b>Strategy II</b>                 |          |          |          |          |
| $T_{R_2}^c$                        | 34.5214  | 31.3491  | 33.2554  | 37.0227  |
| $T_{DP_2}^c / T_{S_2}^c, i = 1, 4$ | 100.8228 | 101.0329 | 100.9761 | 101.1263 |
| $T_{ak_j}^c, j = 2, 8$             | 101.1845 | 101.3445 | 101.3153 | 101.5174 |
| $T_{ak_s}^c$                       | 101.2314 | 101.3935 | 101.3649 | 101.574  |
| <b>Strategy III</b>                |          |          |          |          |
| $T_{R_3}^c$                        | 50.0427  | 46.4561  | 48.6303  | 52.7625  |
| $T_{DP_3}^c / T_{S_3}^c, i = 3, 6$ | 100.4313 | 100.541  | 100.5114 | 100.5896 |
| $T_{ak_j}^c, j = 3, 9$             | 100.7931 | 100.8526 | 100.8506 | 100.9808 |
| $T_{ak_s}^c$                       | 100.8173 | 100.8778 | 100.8761 | 101.0099 |
| $x^* \sim \chi^2(7)$               |          |          |          |          |
| $y^* \sim \chi^2(11)$              |          |          |          |          |
| <b>Strategy I</b>                  |          |          |          |          |
| $T_{R_1}^c$                        | 37.7957  | 36.9291  | 38.9137  | 40.5873  |
| $T_{DP_1}^c / T_{S_1}^c, i = 1, 4$ | 100.1268 | 100.3388 | 100.3466 | 100.8760 |
| $T_{ak_j}^c, j = 1, 7$             | 100.9514 | 101.1073 | 101.1787 | 101.7163 |
| $T_{ak_s}^c$                       | 100.9896 | 101.1671 | 101.2414 | 101.8149 |
| <b>Strategy II</b>                 |          |          |          |          |
| $T_{R_2}^c$                        | 22.3490  | 21.7129  | 23.1803  | 24.4481  |
| $T_{DP_2}^c / T_{S_2}^c, i = 1, 4$ | 100.2680 | 100.7179 | 100.7345 | 101.8676 |
| $T_{ak_j}^c, j = 2, 8$             | 101.0927 | 101.4864 | 101.5666 | 102.7079 |
| $T_{ak_s}^c$                       | 101.1743 | 101.6157 | 101.7022 | 102.9248 |
| <b>Strategy III</b>                |          |          |          |          |
| $T_{R_3}^c$                        | 35.3523  | 34.5106  | 36.4404  | 38.0739  |
| $T_{DP_3}^c / T_{S_3}^c, i = 3, 6$ | 100.1409 | 100.3765 | 100.3852 | 100.9743 |
| $T_{ak_j}^c, j = 3, 9$             | 100.9655 | 101.1451 | 101.2173 | 101.8146 |
| $T_{ak_s}^c$                       | 101.008  | 101.2117 | 101.2872 | 101.9246 |

stratum to stratum, given the sample in every stratum is to be enough big so that the approximate expression of  $MSE(T_{ak_j}^s), j = 1, 2, \dots, 9$  is valid.

### 5. Simulation study

To stimulate the soundness of the theoretical findings, motivated by [31], [32], and [33], a simulation study is presented by utilizing three artificial populations, namely, normal, exponential, and chi-square, each consisting of  $N = 300$  units having the variables  $x$  and  $y$  which can be obtained as

$$y_i = 8.2 + \sqrt{1 - \rho_{xy}^2} y_i^* + \rho_{xy} \left( \frac{S_y}{S_x} \right) x_i^*$$

**Table 2**  
*PRE* of suggested separate estimators.

| Correlation coefficient            | 0.6      | 0.7      | 0.8      | 0.9      |
|------------------------------------|----------|----------|----------|----------|
| $T_m$                              | 100      | 100      | 100      | 100      |
| $x^* \sim N(2, 7)$                 |          |          |          |          |
| $y^* \sim N(4, 9)$                 |          |          |          |          |
| Strategy I                         |          |          |          |          |
| $T_{R_1}^s$                        | 47.4202  | 48.2471  | 48.4044  | 50.0764  |
| $T_{DP_i}^s / T_{S_j}^s, i = 1, 4$ | 100.6625 | 100.8165 | 100.6375 | 100.5616 |
| $T_{ak_j}^s, j = 1, 7$             | 107.5355 | 107.7830 | 107.7864 | 108.2394 |
| $T_{ak_4}^s$                       | 108.1420 | 108.4567 | 108.3925 | 108.8289 |
| Strategy II                        |          |          |          |          |
| $T_{R_2}^s$                        | 29.9329  | 30.6324  | 30.7665  | 32.2095  |
| $T_{DP_i}^s / T_{S_j}^s, i = 1, 4$ | 101.4089 | 101.7394 | 101.3554 | 101.1931 |
| $T_{ak_j}^s, j = 2, 8$             | 108.282  | 108.7059 | 108.5043 | 108.8708 |
| $T_{ak_5}^s$                       | 109.6051 | 110.1820 | 109.8252 | 110.1515 |
| Strategy III                       |          |          |          |          |
| $T_{R_3}^s$                        | 44.8028  | 45.6236  | 45.7799  | 47.4446  |
| $T_{DP_i}^s / T_{S_j}^s, i = 3, 6$ | 100.7366 | 100.908  | 100.7088 | 100.6244 |
| $T_{ak_j}^s, j = 3, 9$             | 107.6097 | 107.8745 | 107.8577 | 108.3022 |
| $T_{ak_6}^s$                       | 108.2858 | 108.6259 | 108.5333 | 108.9590 |
| $x^* \sim Exp(1.0)$                |          |          |          |          |
| $y^* \sim Exp(0.5)$                |          |          |          |          |
| Strategy I                         |          |          |          |          |
| $T_{R_1}^s$                        | 49.3701  | 45.5457  | 47.9883  | 52.5620  |
| $T_{DP_i}^s / T_{S_j}^s, i = 1, 4$ | 100.3880 | 100.4866 | 100.4600 | 100.5303 |
| $T_{ak_j}^s, j = 1, 7$             | 100.7141 | 100.7651 | 100.7663 | 100.8899 |
| $T_{ak_4}^s$                       | 100.7351 | 100.7868 | 100.7884 | 100.9153 |
| Strategy II                        |          |          |          |          |
| $T_{R_2}^s$                        | 31.5957  | 28.3765  | 30.4126  | 34.4198  |
| $T_{DP_i}^s / T_{S_j}^s, i = 1, 4$ | 100.8228 | 101.0329 | 100.9761 | 101.1263 |
| $T_{ak_j}^s, j = 2, 8$             | 101.1488 | 101.3113 | 101.2824 | 101.4858 |
| $T_{ak_5}^s$                       | 101.1941 | 101.3585 | 101.3302 | 101.5409 |
| Strategy III                       |          |          |          |          |
| $T_{R_3}^s$                        | 46.7407  | 42.9472  | 45.3665  | 49.9303  |
| $T_{DP_i}^s / T_{S_j}^s, i = 3, 6$ | 100.4313 | 100.5410 | 100.5114 | 100.5896 |
| $T_{ak_j}^s, j = 3, 9$             | 100.7574 | 100.8194 | 100.8177 | 100.9492 |
| $T_{ak_6}^s$                       | 100.7808 | 100.8436 | 100.8423 | 100.9775 |
| $x^* \sim \chi^2(7)$               |          |          |          |          |
| $y^* \sim \chi^2(11)$              |          |          |          |          |
| Strategy I                         |          |          |          |          |
| $T_{R_1}^s$                        | 34.4668  | 33.5150  | 34.8819  | 36.1917  |
| $T_{DP_i}^s / T_{S_j}^s, i = 1, 4$ | 100.1268 | 100.3388 | 100.3466 | 100.876  |
| $T_{ak_j}^s, j = 1, 7$             | 100.8635 | 101.0234 | 101.0715 | 101.6031 |
| $T_{ak_4}^s$                       | 100.9001 | 101.0808 | 101.1312 | 101.6970 |
| Strategy II                        |          |          |          |          |
| $T_{R_2}^s$                        | 19.9444  | 19.2757  | 20.2386  | 21.1774  |
| $T_{DP_i}^s / T_{S_j}^s, i = 1, 4$ | 100.2680 | 100.7179 | 100.7345 | 101.8676 |
| $T_{ak_j}^s, j = 2, 8$             | 101.0047 | 101.4025 | 101.4595 | 102.5947 |
| $T_{ak_5}^s$                       | 101.0832 | 101.5270 | 101.5887 | 102.8017 |
| Strategy III                       |          |          |          |          |
| $T_{R_3}^s$                        | 32.1275  | 31.2095  | 32.5284  | 33.7957  |
| $T_{DP_i}^s / T_{S_j}^s, i = 3, 6$ | 100.1409 | 100.3765 | 100.3852 | 100.9743 |
| $T_{ak_j}^s, j = 3, 9$             | 100.8776 | 101.0612 | 101.1102 | 101.7014 |
| $T_{ak_6}^s$                       | 100.9184 | 101.1252 | 101.1767 | 101.8062 |

$$x_i = 4.2 + x_i^*$$

where  $x_i^*$  and  $y_i^*$  are independent proportional distribution variables. Following the division of each population into three equal strata, a random sample of size 10 units is taken from each stratum. The suggested combined and separate estimators' percent relative efficiency (*PRE*) regarding the traditional mean estimator is derived after 15000 iterations as

$$PRE = \frac{\frac{1}{15,000} \sum_{i=1}^{15,000} (T_m - \bar{Y})^2}{\frac{1}{15,000} \sum_{i=1}^{15,000} (T - \bar{Y})^2} \times 100$$

where  $T$  denotes the conventional and suggested sequent combined and separate estimators.

The findings of the simulation are unfolded herein from Table 1 to Table 2 in terms of *PRE* for reasonably chosen amounts of the correlation coefficients 0.6, 0.7, 0.8, 0.9 and non responding unit  $r = 5$ .

From Table 1 and Table 2 based on the results of each population, we observe that the suggested CSIMs  $y_{ak_j}^c$  and  $y_{ak_j}^s, j = 1, 2, \dots, 9$  are the most efficient than the other existing IMs for different values of correlation coefficient  $\rho_{xyh}$  under each strategy. It is also seen that the suggested CSIMs  $y_{ak_j}^c$  and  $y_{ak_j}^s, j = 4, 5, 6$  are superior among the suggested CSIMs, respectively.

### 5.1. Results and discussion

From the theoretical and simulation results, we have drawn the following observations:

- (i). The suggested CSIMs  $y_{ak_j}^c$  and  $y_{ak_j}^s, j = 1, 2, \dots, 9$  dominate combined and separate mean IM  $y_{im}^c$  and  $y_{im}^s$ , classical ratio IM  $y_{R_j}^c$  and  $y_{R_j}^s, j = 1, 2, 3$ , ratio kind of IMs  $y_{iS_j}^c$  and  $y_{iS_j}^s, j = 1, 2, \dots, 6$  and regression IMs  $y_{iDP_j}^c$  and  $y_{iDP_j}^s, j = 1, 2, 3$  for different values of correlation coefficient  $\rho_{xyh}$  under strategies I to III.
- (ii). The most effective methods in the recommended class of CSIMs are  $y_{ak_j}^c$  and  $y_{ak_j}^s, j = 4, 5, 6$  under strategies I to III, respectively.
- (iii). The suggested combined IMs  $y_{ak_j}^c, j = 1, 2, \dots, 9$  outperform the suggested separate IMs  $y_{ak_j}^s, j = 1, 2, \dots, 9$  under each strategy in each population.
- (iv). The suggested CSIMs  $y_{ak_j}^c$  and  $y_{ak_j}^s, j = 1, 2, \dots, 9$  perform better under strategy II as compare to strategies I and III.
- (v). The *PRE* of the suggested CSIMs  $y_{ak_j}^c$  and  $y_{ak_j}^s, j = 1, 2, \dots, 9$  increases with the successive incremental change in the value of correlation coefficient from 0.6 to 0.9.

### 6. Illustration

The illustration of the suggested and traditional IMs is demonstrated utilizing real data from [34], page 228. The production of  $N = 80$  factories is taken as the main variable  $y$  and the fixed capital of the factories is taken as the auxiliary variable  $x$ . These variables are noted from four areas (strata) of the 80 factories. Neyman allocation is implemented to select a total sample of size  $n = 45$  from  $h = 4$  strata. The descriptives are given in Table 3.

Based on the descriptives given in Table 3, we have computed the *PRE* of combined and separate estimators  $T$  with respect to the usual mean estimator  $T_m$  utilizing the following formula:

$$PRE = \frac{MSE(T_m)}{MSE(T)} \times 100$$

Table 4 displays the outcomes produced utilizing real data, demonstrating the supremacy of the suggested estimators over the conventional estimators. Moreover, the combined class of estimators perform better than the separate class of estimators in each strategy.

### 7. Conclusion

This study considers some novel classes of CSIMs in the case of missing data in SSRS. The characteristics of the sequent combined and separate estimators obtained from the respective IMs are reported. The members of the suggested combined and separate classes of IMs are identified as the combined and separate conventional mean IM, ratio IM, ratio type IM defined on the lines of [27], and regression IM defined on the lines of [4]. The theoretical conditions are derived and sustained with the simulation study accomplished on the hypothetically drawn one symmetrical population such as normal and two asymmetrical populations such as exponential and Chi-square. The simulation



**Table 3**  
Descriptive statistics of real data.

|                         | Total                          | Symbol for stratum <i>h</i>      | 1       | 2       | 3       | 4       |
|-------------------------|--------------------------------|----------------------------------|---------|---------|---------|---------|
| Population size         | <i>N</i> = 80                  | <i>N<sub>h</sub></i>             | 19      | 32      | 14      | 15      |
| Sample size             | <i>n</i> = 45                  | <i>n<sub>h</sub></i>             | 11      | 18      | 8       | 8       |
| Responding units        | <i>r</i> = 26                  | <i>r<sub>h</sub></i>             | 6       | 12      | 4       | 4       |
| Population mean         | $\bar{X} = 1126.46$            | $\bar{X}_h$                      | 349.68  | 706.59  | 1539.57 | 2620.53 |
| Population mean         | $\bar{Y} = 5182.64$            | $\bar{Y}_h$                      | 2967.95 | 4657.63 | 6537.21 | 7843.67 |
| Correlation coefficient | $\rho_{xy} = 0.94$             | $\rho_{xy_h}$                    | 0.93    | 0.92    | 0.98    | 0.96    |
| Standard deviation      | <i>S<sub>x</sub></i> = 845.61  | <i>S<sub>x<sub>h</sub></sub></i> | 109.44  | 109.22  | 277.18  | 370.97  |
| Standard deviation      | <i>S<sub>y</sub></i> = 1835.66 | <i>S<sub>y<sub>h</sub></sub></i> | 757.08  | 669.12  | 416.11  | 645.68  |

**Table 4**  
*PRE* of suggested combined estimators w.r.t. usual mean estimator.

| Combined estimators                | PRE       | Separate estimators                | PRE       |
|------------------------------------|-----------|------------------------------------|-----------|
| <b>Strategy I</b>                  |           |                                    |           |
| $T_{R_1}^c$                        | 113.5945  | $T_{R_1}^s$                        | 112.9952  |
| $T_{DP_1}^c / T_{S_1}^c, i = 1, 4$ | 142.6297  | $T_{DP_1}^s / T_{S_1}^s, i = 1, 4$ | 141.1211  |
| $T_{ak_j}^c, j = 1, 7$             | 143.0664  | $T_{ak_j}^s, j = 1, 7$             | 142.7539  |
| $T_{ak_4}^c$                       | 143.2226  | $T_{ak_4}^s$                       | 142.7983  |
| <b>Strategy II</b>                 |           |                                    |           |
| $T_{R_2}^c$                        | 160.2879  | $T_{R_2}^s$                        | 159.8929  |
| $T_{DP_2}^c / T_{S_1}^c, i = 2, 5$ | 1648.7680 | $T_{DP_2}^s / T_{S_1}^s, i = 2, 5$ | 1647.9714 |
| $T_{ak_j}^c, j = 2, 8$             | 1649.2040 | $T_{ak_j}^s, j = 2, 8$             | 1648.8920 |
| $T_{ak_5}^c$                       | 1670.3520 | $T_{ak_5}^s$                       | 1654.8650 |
| <b>Strategy III</b>                |           |                                    |           |
| $T_{R_3}^c$                        | 134.4895  | $T_{R_3}^s$                        | 133.1935  |
| $T_{DP_3}^c / T_{S_1}^c, i = 3, 6$ | 278.1370  | $T_{DP_3}^s / T_{S_1}^s, i = 3, 6$ | 277.7919  |
| $T_{ak_j}^c, j = 3, 9$             | 278.5738  | $T_{ak_j}^s, j = 3, 9$             | 278.2613  |
| $T_{ak_6}^c$                       | 279.3849  | $T_{ak_6}^s$                       | 278.4918  |

results exhibit the dominance of the suggested CSIMs over the existing CSIMs. The numerical findings based on real data also exhibit the superiority of the proposed IMs over their traditional counterparts. Furthermore, the combined IMs perform superior than the separate IMs. As a result, it is necessary to favor the suggested combined and separate classes of IMs for estimating the population mean in case of missing data.

The suggested CSIMs will be evaluated under SRSS in upcoming studies.

**Declaration of competing interest**

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this article.

**Data availability**

The data chosen for the empirical study are provided in the manuscript.

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**Appendix A**

The MSE of the subsequent estimators used in currently available combined imputation approaches is displayed below.

$$V(T_m) = \bar{Y}^2 I_0^* \tag{A.1}$$

$$MSE(T_{R_1}^c) = \bar{Y}^2 [I_0^* + I_1 - 2I_{01}] \tag{A.2}$$

$$MSE(T_{R_2}^c) = \bar{Y}^2 [I_0^* + I_1^* - 2I_{01}^*] \tag{A.3}$$

$$MSE(T_{R_3}^c) = \bar{Y}^2 [I_0^* + I_1^* - I_1 - 2(I_{01}^* - I_{01})] \tag{A.4}$$

$$MSE(T_{DP_1}^c) = \bar{Y}^2 [I_0^* + b_1^2 \bar{X}^2 I_1 - 2b_1 \bar{X} \bar{Y} I_{01}] \tag{A.5}$$

$$MSE(T_{DP_2}^c) = \bar{Y}^2 [I_0^* + b_2^2 \bar{X}^2 I_1^* - 2b_2 \bar{X} \bar{Y} I_{01}^*] \tag{A.6}$$

$$MSE(T_{DP_3}^c) = \bar{Y}^2 [I_0^* + b_3^2 \bar{X}^2 (I_1^* - I_1) - 2b_3 \bar{X} \bar{Y} (I_{01}^* - I_{01})] \tag{A.7}$$

$$\min MSE(T_{DP_1}^c) = \bar{Y}^2 \left[ I_0^* - \frac{I_{01}^2}{I_1} \right] \tag{A.5}$$

$$\min MSE(T_{DP_2}^c) = \bar{Y}^2 \left[ I_0^* - \frac{I_{01}^{*2}}{I_1^*} \right] \tag{A.6}$$

$$\min MSE(T_{DP_3}^c) = \bar{Y}^2 \left[ I_0^* - \frac{(I_{01}^* - I_{01})^2}{(I_1^* - I_1)} \right] \tag{A.7}$$

$$MSE(T_{S_i}^c) = \bar{Y}^2 [I_0^* + \beta_i^2 I_1 - 2\beta_i I_{01}], i = 1, 4 \tag{A.8}$$

$$MSE(T_{S_i}^c) = \bar{Y}^2 [I_0^* + \beta_i^2 I_1^* - 2\beta_i I_{01}^*], i = 2, 5 \tag{A.9}$$

$$MSE(T_{S_i}^c) = \bar{Y}^2 \left[ I_0^* + \beta_i^2 \left\{ \frac{I_1^* - I_1}{-2\beta_i} \left\{ \frac{I_{01}^* - I_{01}}{I_1^* - I_1} \right\} \right\} \right], i = 3, 6 \tag{A.10}$$

$$\min MSE(T_{S_i}^c) = \bar{Y}^2 \left[ I_0^* - \frac{I_{01}^2}{I_1} \right]; i = 1, 4 \tag{A.8}$$

$$\min MSE(T_{S_i}^c) = \bar{Y}^2 \left[ I_0^* - \frac{I_{01}^{*2}}{I_1^*} \right]; i = 2, 5 \tag{A.9}$$

$$\min MSE(T_{S_i}^c) = \bar{Y}^2 \left[ I_0^* - \frac{(I_{01}^* - I_{01})^2}{(I_1^* - I_1)} \right]; i = 3, 6 \tag{A.10}$$

The scalars' optimum values for the combined estimators are shown below.

$$b_1 = R \frac{I_{01}}{I_1}, b_2 = R \frac{I_{01}^*}{I_1^*}, b_3 = R \frac{(I_{01}^* - I_{01})}{(I_1^* - I_1)}, \beta_{1(opt)} = \beta_{4(opt)} = \frac{I_{01}}{I_1}, \beta_{2(opt)} = \frac{I_{01}^*}{I_1^*}, \text{ and } \beta_{3(opt)} = \beta_{6(opt)} = \frac{(I_{01}^* - I_{01})}{(I_1^* - I_1)}.$$

**Appendix B**

The MSE of the subsequent estimators used in currently available separate imputation approaches is displayed below.

$$MSE(T_{R_1}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 (J_1 + J_0^* - 2J_{01}) \tag{B.11}$$

$$MSE(T_{R_2}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 (J_1^* + J_0^* - 2J_{01}^*) \tag{B.12}$$

$$MSE(T_{R_3}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 [J_0^* + (J_1^* - J_1) - 2(J_{01}^* - J_{01})] \tag{B.13}$$

$$MSE(T_{DP_1}^s) = \sum_{h=1}^L W_h^2 [\bar{Y}_h^2 J_0^* + b_{1_h}^2 \bar{X}_h^2 J_1 - 2b_{1_h} \bar{X}_h \bar{Y}_h J_{01}]$$

$$MSE(T_{DP_2}^s) = \sum_{h=1}^L W_h^2 [\bar{Y}_h^2 J_0^* + b_{2_h}^2 \bar{X}_h^2 J_1^* - 2b_{2_h} \bar{X}_h \bar{Y}_h J_{01}^*]$$

$$MSE(T_{DP_3}^s) = \sum_{h=1}^L W_h^2 [\bar{Y}_h^2 J_0^* + b_{3_h}^2 \bar{X}_h^2 (J_1^* - J_1) - 2b_{3_h} \bar{X}_h \bar{Y}_h (J_{01}^* - J_{01})]$$

$$\min MSE(T_{DP_1}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{J_{01}^2}{J_1} \right] \tag{B.14}$$

$$\min MSE(T_{DP_2}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{J_{01}^2}{J_1^*} \right] \tag{B.15}$$

$$\min MSE(T_{DP_3}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{(J_{01}^* - J_{01})^2}{(J_1^* - J_1)} \right] \tag{B.16}$$

$$MSE(T_{S_i}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 [J_0^* + \beta_{i_h}^2 J_1 - 2\beta_{i_h} J_{01}], \quad i = 1, 4$$

$$MSE(T_{S_i}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 [J_0^* + \beta_{i_h}^2 J_1^* - 2\beta_{i_h} J_{01}^*], \quad i = 2, 5$$

$$MSE(T_{S_i}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 [J_0^* + \beta_{i_h}^2 (J_1^* - J_1) - 2\beta_{i_h} (J_{01}^* - J_{01})],$$

$i = 3, 6$

$$\min MSE(T_{S_i}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{J_{01}^2}{J_1} \right]; \quad i = 1, 4 \tag{B.17}$$

$$\min MSE(T_{S_i}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{J_{01}^2}{J_1^*} \right]; \quad i = 2, 5 \tag{B.18}$$

$$\min MSE(T_{S_i}^s) = \sum_{h=1}^L W_h^2 \bar{Y}_h^2 \left[ J_0^* - \frac{(J_{01}^* - J_{01})^2}{(J_1^* - J_1)} \right]; \quad i = 3, 6 \tag{B.19}$$

The scalars' optimum values for the combined estimators are shown below.

$$b_{1_h} = R_h \frac{J_{01}}{J_1}, \quad b_{2_h} = R_h \frac{J_{01}^*}{J_1^*}, \quad b_{3_h} = R_h \frac{(J_{01}^* - J_{01})}{(J_1^* - J_1)}, \quad \beta_{1_h(opt)} = \beta_{4_h(opt)} = \frac{J_{01}}{J_1},$$

$$\beta_{2_h(opt)} = \beta_{5_h(opt)} = \frac{J_{01}^*}{J_1^*}, \quad \beta_{3_h(opt)} = \beta_{6_h(opt)} = \frac{(J_{01}^* - J_{01})}{(J_1^* - J_1)}.$$

### Appendix C

This section contains the proof of MSE and minimum MSE expressions discussed in Section 3.1.1 and Section 3.2.1.

Consider the following estimator as part of Strategy I:

$$T_{ak_1}^c = \alpha_1 \bar{y}_{rst} + \theta_1 (\bar{x}_{nst} - \bar{X})$$

Employing the notations discussed in subsection 2.1, we get

$$T_{ak_1}^c - \bar{Y} = \alpha_1 \bar{Y} \varepsilon_0 + \theta_1 \bar{X} \varepsilon_1 + (\alpha_1 - 1) \bar{Y} \tag{C.20}$$

By squaring both sides of (C.20) and taking the expectation, we get

$$MSE(T_{ak_1}^c) = (\alpha_1 - 1)^2 \bar{Y}^2 + \alpha_1^2 \bar{Y}^2 I_0^* + \theta_1^2 I_1 + 2\alpha_1 \theta_1 \bar{X} \bar{Y} I_{01} \tag{C.21}$$

By minimizing (C.21) w.r.t  $\alpha_1$  and  $\theta_1$ , we get the optimum values of  $\alpha_1$  and  $\theta_1$  as

$$\alpha_{1(opt)} = \frac{1}{\left(1 + I_0^* - \frac{I_{01}^2}{I_1}\right)} = \alpha_{7(opt)}$$

$$\theta_{1(opt)} = -\frac{\bar{Y}}{\bar{X}} \frac{I_{01}}{I_1} \alpha_{1(opt)}$$

Putting  $\alpha_{1(opt)}$  and  $\theta_{1(opt)}$  in (C.21), we get the minimum MSE as

$$MSE_{min}(T_{ak_1}^c) = \bar{Y}^2 (1 - \alpha_{1(opt)})$$

In a similar manner, the minimum MSE of additional suggested estimators can be determined. The scalars' optimum values for the suggested combined estimators are shown below.

$$\alpha_{2(opt)} = \frac{1}{\left(1 + I_0^* - \frac{I_{01}^2}{I_1^*}\right)} = \alpha_{8(opt)}$$

$$\theta_{2(opt)} = -\frac{\bar{Y}}{\bar{X}} \frac{I_{01}^*}{I_1^*} \alpha_{2(opt)}$$

$$\alpha_{3(opt)} = \frac{1}{\left(1 + I_0^* - \frac{(I_{01}^* - I_{01})^2}{(I_1^* - I_1)}\right)}$$

$$\theta_{3(opt)} = -\frac{\bar{Y}}{\bar{X}} \left( \frac{I_{01}^* - I_{01}}{I_1^* - I_1} \right) \alpha_{3(opt)}$$

$$\alpha_{j(opt)} = \frac{A_j}{B_j}; \quad j = 4, 5, 6, 7, 8, 9$$

$$\theta_{j(opt)} = \frac{I_{01}}{I_1}; \quad j = 4, 7$$

$$\theta_{j(opt)} = \frac{I_{01}^*}{I_1^*}; \quad j = 5, 8.$$

$$\theta_{j(opt)} = \frac{(I_{01}^* - I_{01})}{(I_1^* - I_1)}; \quad j = 6, 9$$

where  $A_4 = \left(1 + \frac{I_{01}}{2} - \frac{I_{12}^2}{2I_1}\right)$ ,  $B_4 = \left(1 + I_0^* + I_{01} - \frac{2I_{12}^2}{I_1}\right)$ ,  $A_5 = \left(1 + \frac{I_{01}^*}{2} - \frac{I_{01}^{*2}}{2I_1^*}\right)$ ,  $B_5 = \left(1 + I_0^* + I_{01}^* - \frac{2I_{01}^{*2}}{I_1^*}\right)$ ,  $A_6 = \left[1 - \frac{1}{2} \frac{\{I_{01}^* - I_{01}\}^2}{I_1^* - I_1} + \frac{1}{2} (I_{01}^* - I_{01})\right]$ ,  $B_6 = \left[1 + I_0^* - 2 \frac{\{I_{01}^* - I_{01}\}^2}{I_1^* - I_1} + (I_{01}^* - I_{01})\right]$ ,  $A_7 = 1$ ,  $B_7 = \left(1 + I_0^* - \frac{I_{12}^2}{I_1}\right)$ ,  $A_9 = 1$  and  $B_9 = \left[1 + I_0^* - \frac{(I_{01}^* - I_{01})^2}{(I_1^* - I_1)}\right]$ .

Similarly, the outline of the proof of MSE and minimum MSE expressions given in Section 3.2.1 can be obtained.

### References

- [1] R.B. Rubin, Inference and missing data, *Biometrika* 63 (3) (1976) 581–592.
- [2] D.F. Heitjan, S. Basu, Distinguishing ‘missing at random’ and ‘missing completely at random’, *Am. Stat.* 50 (1996) 207–213.
- [3] C. Kadilar, H. Cingi, Estimators for the population mean in the case of missing data, *Commun. Stat., Theory Methods* 37 (2008) 2226–2236.
- [4] G. Diana, P.F. Perri, Improved estimators of the population mean for missing data, *Commun. Stat., Theory Methods* 39 (2010) 3245–3251.
- [5] S. Bhushan, A.P. Pandey, Optimal imputation of missing data for estimation of population mean, *J. Stat. Manag. Syst.* 19 (6) (2016) 755–769.
- [6] C. Mohamed, S.A. Sedory, S. Singh, Imputation using higher order moments of an auxiliary variable, *Commun. Stat., Simul. Comput.* 46 (2016) 6588–6617.
- [7] S. Bhushan, A.P. Pandey, A. Pandey, On optimality of imputation methods for estimation of population mean using higher order moment of an auxiliary variable, *Commun. Stat., Simul. Comput.* 49 (6) (2020) 1560–1574.
- [8] G.N. Singh, M. Usman, Generalized class of estimators using dual of auxiliary variable under non-response, *J. Stat. Theory Pract.* 15 (2) (2021) 1–35.
- [9] M.M. Anas, Z. Huang, U. Shahzad, T. Zaman, S. Shahzadi, Compromised imputation based mean estimators using robust quantile regression, *Commun. Stat., Theory Methods* (2022) 1–16.
- [10] S. Bhushan, A. Kumar, A.P. Pandey, S. Singh, Estimation of population mean in presence of missing data under simple random sampling, *Commun. Stat., Simul. Comput.* 52 (12) (2023) 6048–6069.
- [11] S. Bhushan, A. Kumar, T. Zaman, A. Al Mutairi, Efficient difference and ratio-type imputation methods under ranked set sampling, *Axioms* 12 (2023) 558.
- [12] S. Bhushan, A. Kumar, Imputation of missing data using multi auxiliary information under ranked set sampling, *Commun. Stat., Simul. Comput.* (2023) 1–23, <https://doi.org/10.1080/03610918.2023.2288796>.

- [13] U. Shahzad, M. Hanif, N. Koyuncu, A new estimator for mean under stratified random sampling, *Math. Sci.* (2018) 163–169.
- [14] T. Zaman, C. Kadilar, On estimating the population mean using auxiliary character in stratified random sampling, *J. Stat. Manag. Syst.* 23 (8) (2020) 1415–1426.
- [15] T. Zaman, An efficient exponential estimator of the mean under stratified random sampling, *Math. Popul. Stud.* 28 (2) (2021) 104–121.
- [16] M. Subzar, S.A. Lone, M. Aslam, A.H. AL-Marshadi, S. Maqbool, Exponential ratio estimator of the median: an alternative to the regression estimator of the median under stratified sampling, *J. King Saud Univ., Sci.* 35 (3) (2023) 102536.
- [17] S. Ahmad, S. Hussain, M. Aamir, U. Yasmeen, J. Shabbir, Z. Ahmad, Dual use of auxiliary information for estimating the finite population mean under the stratified random sampling scheme, *J. Math.* 2021 (2021) 1–2.
- [18] S. Ahmad, S. Hussain, U. Yasmeen, M. Aamir, J. Shabbir, M. El-Morshedy, A. Al-Bossly, Z. Ahmad, A simulation study: using dual ancillary variable to estimate population mean under stratified random sampling, *PLoS ONE* 17 (11) (2022) e0275875.
- [19] T. Zaman, C. Kadilar, Exponential ratio and product type estimators of the mean in stratified two-phase sampling, *AIMS Math.* 6 (5) (2021) 4265–4279.
- [20] Q. Rana, M.N. Qureshi, M. Hanif, Generalized estimator for population mean using auxiliary attribute in stratified two-phase sampling, *J. Stat. Theory Appl.* 21 (2) (2022) 44–57.
- [21] S. Bhushan, A. Kumar, S. Singh, Some efficient classes of estimators under stratified sampling, *Commun. Stat., Theory Methods* (2021) 1–30, <https://doi.org/10.1080/03610926.2021.1939052>.
- [22] S. Bhushan, A. Kumar, R. Onyango, S. Singh, Some improved classes of estimators in stratified sampling using bivariate auxiliary information, *J. Probab. Stat.* 2022 (2) (2022) 1–23.
- [23] S. Bhushan, A. Kumar, U. Shahzad, A.I. Al-Omari, A.I. Almanjahie, On some improved class of estimators by using stratified ranked set sampling, *Mathematics* 10 (3283) (2022) 1–32.
- [24] S. Bhushan, A. Kumar, S. Shahab, S.A. Lone, M.T. Akhtar, On efficient estimation of population mean under stratified ranked set sampling, *J. Math.* 2022 (3) (2022) 1–20.
- [25] S. Bhushan, A. Kumar, J. Banerjee, Mean estimation using logarithmic estimators in stratified ranked set sampling, *Life Cycle Reliab. Saf. Eng.* (2022) 1–9.
- [26] N.S. Thakur, K. Yadav, S. Pathak, Estimation of mean with imputation of missing data in stratified random sampling, *Res. Rev.: J. Stat.* 3 (1) (2014) 1–12.
- [27] M.S. Ahmed, O. Al-Titi, Z. Al-Rawi, W. Abu-Dayyeh, Estimation of a population mean using different imputation methods, *Stat. Transit.* 7 (6) (2006) 1247–1264.
- [28] R. Pandey, N.S. Thakur, K. Yadav, Separate regression type imputation methods to estimate population mean, *Int. J. Eng. Technol. Sci. Res.* 2 (2015) 36–45.
- [29] S. Bhushan, A.P. Pandey, Optimality of ratio type estimation methods for population mean in the presence of missing data, *Commun. Stat., Theory Methods* 47 (11) (2018) 2576–2589.
- [30] S. Bhushan, A. Kumar, On optimal classes of estimators under ranked set sampling, *Commun. Stat., Theory Methods* 51 (8) (2022) 2610–2639.
- [31] S. Singh, S. Horn, An alternative survey in multi-character survey, *Metrika* 48 (1998) 99–107.
- [32] S. Bhushan, A. Kumar, M.T. Akhtar, S.A. Lone, Logarithmic type predictive estimators under simple random sampling, *AIMS Math.* 7 (7) (2022) 11992–12010.
- [33] S. Bhushan, A. Kumar, Predictive estimation approach using difference and ratio type estimators in ranked set sampling, *J. Comput. Appl. Math.* 410 (2022) 114214.
- [34] M.N. Murthy, *Sampling Theory and Methods*, Statistical Publishing Society, Calcutta, India, 1967.